

Package ‘afex’

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Type Package

Title Analysis of Factorial Experiments

Depends R (>= 3.0.0), car, lme4 (>= 1.0.5), pbkrtest (>= 0.3-6), reshape2

Suggests ascii, xtable, parallel, plyr

Imports stringr, coin, Matrix

Description Provides convenience functions for analyzing factorial experiments using ANOVA or mixed-models. `ez.glm()` and `aov.car()` allow convenient calculation of between, within (i.e., repeated-measures), or mixed between-within (i.e., split-plot) ANOVAs for data in the long format (i.e., one observation per row) wrapping `car::Anova()` (aggregating more than one observation per individual and cell of the design), per default returning a print ready ANOVA table. Function `mixed()` fits a mixed model using `lme4::lmer()` and computes p-values for all effects in the model using either the Kenward-Roger approximation of degrees of freedom (LMM only), parametric bootstrap (LMMs and GLMMs) or likelihood ratio tests (LMMs and GLMMs). `afex` uses type 3 sums of squares as default (imitating commercial statistical software). Furthermore, `compare.2.vectors()` conveniently compares two vectors using a variety of tests.

URL <http://www.psychologie.uni-freiburg.de/Members/singmann/R/afex>

License GPL (>= 3)

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Author Henrik Singmann [aut, cre], Ben Bolker [ctb], Søren Højsgaard [ctb], John Fox [ctb], Michael A. Lawrence [ctb], Ulf Mertens [ctb]

Maintainer Henrik Singmann <henrik.singmann@psychologie.uni-freiburg.de>

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Description

Analysis of Factorial Experiments.

Details

```

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Provides convenience functions for analyzing factorial experiments using ANOVA or mixed-models. `ez.glm()` and `aov.car()` allow convenient calculation of between, within (i.e., repeated-measures), or

mixed between-within (i.e., split-plot) ANOVAs for data in the long format (i.e., one observation per row) wrapping `car::Anova()` (aggregating more than one observation per individual and cell of the design), per default returning a print ready ANOVA table. Function `mixed()` fits a mixed model using `lme4::lmer()` and computes p-values for all effects in the model using either the Kenward-Roger approximation of degrees of freedom (LMM only), parametric bootstrap (LMMs and GLMMs) or likelihood ratio tests (LMMs and GLMMs). `afex` uses type 3 sums of squares as default (imitating commercial statistical software). Furthermore, `compare.2.vectors()` conveniently compares two vectors using a variety of tests.

Author(s)

Henrik Singmann, with contributions from Ben Bolker, Søren Højsgaard, John Fox, Michael A. Lawrence, Ulf Mertens

aov.car	<i>Convenience wrappers for car::Anova using either a formula or factor based interface.</i>
---------	--

Description

These functions allow convenient access to [Anova](#) (from the **car** package) for data in the **long** format (i.e., one observation per row), possibly aggregating the data if there is more than one observation per individual and cell. Hence, mixed between-within ANOVAs can be calculated conveniently without using the rather unhandy format of `car::Anova`. `aov.car` can be called using a formula similar to [aov](#) specifying an error strata for the within-subject factor(s), `aov4` can be called with a **lme4**-like formula, and `ez.glm` is called specifying the factors as character vectors.

Usage

```
aov.car(formula, data, fun.aggregate = NULL, type = 3,
        factorize = TRUE, check.contrasts = TRUE,
        return = "nice", observed = NULL, args.return = list(), ...)

aov4(formula, data, observed = NULL, fun.aggregate = NULL, type = 3,
     factorize = TRUE, check.contrasts = TRUE,
     return = "nice", args.return = list(), ...,
     print.formula = FALSE)

ez.glm(id, dv, data, between = NULL, within = NULL, covariate = NULL,
       observed = NULL, fun.aggregate = NULL, type = 3,
       factorize = TRUE, check.contrasts = TRUE,
       return = "nice", args.return = list(), ..., print.formula = FALSE)

univ(object)
```

Arguments

formula	A formula specifying the ANOVA model similar to aov (for <code>aov.car</code> or similar to <code>lme4:lmer</code> for <code>aov4</code>). Should include an error term (i.e., <code>Error(id/)</code> or <code>(... id)</code>). Note that the within-subject factors do not need to be outside the Error term (this contrasts with <code>aov</code>). See Details.
id	character vector (of length 1) indicating the subject identifier column in data.
dv	character vector (of length 1) indicating the column containing the dependent variable in data.
between	character vector indicating the between -subject(s) factor(s)/column(s) in data. Default is NULL indicating no between-subjects factors.
within	character vector indicating the within -subject(s) factor(s)/column(s) in data. Default is NULL indicating no within-subjects factors.
covariate	character vector indicating the between-subject(s) covariate(s) (i.e., column(s)) in data. Default is NULL indicating no covariates.
observed	character vector indicating which of the variables are observed (i.e, measured) as compared to experimentally manipulated. The default behavior is to return a ANOVA table with generalized eta squared effect size for which this information is necessary (see nice.anova).
data	A <code>data.frame</code> containing the data. Mandatory.
fun.aggregate	The function for aggregating the data before running the ANOVA if there is more than one observation per individual and cell of the design. The default NULL issues a warning if aggregation is necessary and uses mean .
type	The type of sums of squares for the ANOVA. Defaults to 3 . Passed to Anova . Possible values are "II", "III", 2, or 3.
factorize	logical. Should between subject factors be factorized (with note) before running the analysis. Default is TRUE. If one wants to run an ANCOVA, needs to be set to FALSE (in which case centering on 0 is checked on numeric variables).
check.contrasts	logical. Should contrasts for between-subject factors be checked and (if necessary) changed to be <code>"contr.sum"</code> . See details.
print.formula	<code>ez.glm</code> is a wrapper for <code>aov.car</code> . This boolean argument indicates whether the formula in the call to <code>car.aov</code> should be printed.
return	What should be returned? If "nice" (the default) will return a nice ANOVA table (produced by nice.anova . Possible values are <code>c("Anova", "lm", "data", "nice", "full", "all</code> (possibly abbreviated).
args.return	list of further arguments passed to the function which produces the return value. Currently only supports <code>return = "nice"</code> (the default) which then passes arguments to nice.anova (see examples).
...	Further arguments passed to <code>fun.aggregate</code> .
object	An object of class <code>Anova.mlm</code> as returned by <code>aov.car</code> , <code>ez.glm</code> , or Anova .

Details

Type 3 sums of squares are default in afex. Note that type 3 sums of squares are said to be dangerous and/or problematic. On the other side they are the default in in SPSS and SAS and recommended by e.g. Maxwell and Delaney (2004). For a brief discussion see [here](#).

However, note that lower order effects (e.g., main effects) in type 3 ANOVAs are only meaningful with **effects coding**. That is, contrasts should be set to `contr.sum` via `options(contrasts=c('contr.sum', 'contr.poly'))`. This should be done automatically when loading **afex** and **afex** will issue a warning when running type 3 SS and **other coding schemes**. You can check the coding with `options("contrasts")`.

The formulas for `aov.car` or `aov4` must contain a single Error term specifying the ID column and potential within-subject factors (you can use `mixed` with multiple error terms). Factors outside the Error term are treated as between-subject factors (the within-subject factors specified in the Error term are ignored outside the Error term, i.e., it is not necessary to specify them outside the Error term, see Examples).

Suppressing the intercept (i.e. via $0 +$ or $- 1$) is ignored. Specific specifications of effects (e.g., excluding terms with $-$ or using $^$) could be okay but is not tested. Using the `I` or `poly` function within the formula is not tested and not supported!

For `ez.glm` either `between` or `within` must not be NULL.

`ez.glm` will concatenate all between-subject factors using `*` (i.e., producing all main effects and interactions) and all covariates by `+` (i.e., adding only the main effects to the existing between-subject factors). The within-subject factors do fully interact with all between-subject factors and covariates. This is essentially identical to the behavior of SPSS's `glm` function.

To run an ANCOVA you need to set `factorize = FALSE` and make sure that all variables have the correct type (i.e., factors are factors and numeric variables are numeric and centered).

Note that the default behavior is to return a `nice.anova` data.frame. This includes calculation of generalized eta squared for which **all non manipulated (i.e., observed)** variables need to be specified via the `observed` argument. Changing the effect size to "pes" (partial eta-squared) via `args.return` or the return value via `return` removes this necessity.

If `check.contrasts = TRUE`, contrasts will be set to "contr.sum" for all between-subject factors if default contrasts are not equal to "contr.sum" or `attrib(factor, "contrasts") != "contr.sum"`. (within-subject factors are hard-coded "contr.sum".)

Value

`aov.car`, `aov4`, and `ez.glm` are wrappers to `Anova`, the return value is dependent on the return argument. When argument `return` is "nice" (the default) a nice ANOVA table is returned (`nice.anova`) with the following columns: Effect, df, MSE (mean-squared errors), F (potentially with significant symbols), ges (generalized eta-squared), p.

If `return = "full"` or `return = "all"` a list with the following elements:

"Anova" the same as `Anova`. Usually an object of class "Anova.mlm" (with within-subjects factors) or of class `c("anova", "data.frame")`. Also returned if `return = "Anova"`.

"lm" the object fitted with `lm` and passed to `Anova` (i.e., an object of class "lm" or "mlm"). Also returned if `return = "lm"`.

"data" the data used to fit the `lm` object. Also returned if `return = "data"`.

"idata" if within-subject factors are present, the `idata` argument passed to `Anova`.

"marginal" a list containing the marginal means (the same as when `return = "marginal"`).

If `return = "univariate"` the object returned from `univ`.

`univ` returns a list of data.frames containing the univariate results (i.e., the classical ANOVA results) from an object of class `"Anova.mlm"`. This is essentially the output from `summary.Anova.mlm` with `multivariate = FALSE`, e.g. `summary(aov.car(...), multivariate = FALSE)`, as a list instead of printed to the console.

For objects of class `"anova"` (i.e., the object returned by `car::Anova` for a purely between-subjects ANOVA) the object is returned unaltered.

The elements of the list returned by `univ` are: `anova`, `mauchly`, and `sphericity.correction` (containing both, Greenhouse-Geisser and Hyundt-Feldt correction).

If `return = "marginal"` A list with data.frames containing the marginal means for all effects. Numerical variables are ignored.

Note

Variables entered as within-subjects (i.e., repeated measures) factors are silently converted to factors and unused levels dropped.

Contrasts attached to a factor as an attribute are probably not preserved and not supported.

The workhorse is `aov.car`. `aov4` and `ez.glm` only construe and pass an appropriate formula to `aov.car`. Use `print.formula = TRUE` to view this formula.

Author(s)

`univ` is basically a copy of `summary.Anova.mlm` written by John Fox.

The other functions were written by Henrik Singmann.

The design of these functions is heavily influenced by `ezANOVA` from package `ez`.

References

Maxwell, S. E., & Delaney, H. D. (2004). *Designing Experiments and Analyzing Data: A Model-Comparisons Perspective*. Mahwah, N.J.: Lawrence Erlbaum Associates.

See Also

`nice.anova` creates the nice ANOVA tables which are by default returned. See also there for a slightly longer discussion of effect sizes.

`mixed` provides a (formula) interface for obtaining p-values for mixed-models via **lme4**.

`obk.long` describes the long version of the OBrienKaiser dataset used in the examples.

Examples

```
# Examples from a pureyl within-design from
# Maxwell & Delaney (2004, Chapter 11),
# Table 12.5 (p. 578):
data(md_12.1)
ez.glm("id", "rt", md_12.1, within = c("angle", "noise"),
```

```

args.return=list(correction = "none", es = "none"))

# Default output
ez.glm("id", "rt", md_12.1, within = c("angle", "noise"))

# examples using obk.long (see ?obk.long), a long version of the OBrienKaiser dataset from car.

data(obk.long, package = "afex")

# run univariate mixed ANOVA for the full design:
aov.car(value ~ treatment * gender + Error(id/phase*hour),
        data = obk.long, observed = "gender")

aov4(value ~ treatment * gender + (phase*hour|id),
      data = obk.long, observed = "gender")

ez.glm("id", "value", obk.long, between = c("treatment", "gender"),
       within = c("phase", "hour"), observed = "gender")

# both calls return the same:
##          Effect          df    MSE      F ges      p
## 1          treatment      2, 10 22.81   3.94 + .20   .05
## 2           gender        1, 10 22.81   3.66 + .11   .08
## 3    treatment:gender      2, 10 22.81   2.86 .18   .10
## 4           phase 1.60, 15.99  5.02 16.13 *** .15 .0003
## 5    treatment:phase 3.20, 15.99  5.02   4.85 * .10   .01
## 6     gender:phase 1.60, 15.99  5.02   0.28 .003   .71
## 7    treatment:gender:phase 3.20, 15.99  5.02   0.64 .01   .61
## 8            hour 1.84, 18.41  3.39 16.69 *** .13 <.0001
## 9    treatment:hour 3.68, 18.41  3.39   0.09 .002   .98
## 10   gender:hour 1.84, 18.41  3.39   0.45 .004   .63
## 11   treatment:gender:hour 3.68, 18.41  3.39   0.62 .01   .64
## 12     phase:hour 3.60, 35.96  2.67   1.18 .02   .33
## 13   treatment:phase:hour 7.19, 35.96  2.67   0.35 .009   .93
## 14   gender:phase:hour 3.60, 35.96  2.67   0.93 .01   .45
## 15  treatment:gender:phase:hour 7.19, 35.96  2.67   0.74 .02   .65

# replicating ?Anova using aov.car:
aov.car(value ~ treatment * gender + Error(id/phase*hour),
        data = obk.long, type = 2, return = "Anova")
# in contrast to aov you do not need the within-subject factors outside Error()

# replicating ?Anova using ez.glm:
ez.glm("id", "value", obk.long, c("treatment", "gender"),
       c("phase", "hour"), type = 2, return = "Anova")

#both return:
## Type II Repeated Measures MANOVA Tests: Pillai test statistic
##          Df test stat approx F num Df den Df      Pr(>F)
## (Intercept)      1    0.970      318      1    10 0.0000000065 ***
## treatment        2    0.481        5      2     10    0.03769 *

```

```

## gender          1    0.204      3    1    10    0.14097
## treatment:gender  2    0.364      3    2    10    0.10447
## phase           1    0.851     26    2     9    0.00019 ***
## treatment:phase  2    0.685      3    4    20    0.06674 .
## gender:phase     1    0.043      0    2     9    0.82000
## treatment:gender:phase  2    0.311      1    4    20    0.47215
## hour            1    0.935     25    4     7    0.00030 ***
## treatment:hour   2    0.301      0    8    16    0.92952
## gender:hour      1    0.293      1    4     7    0.60237
## treatment:gender:hour  2    0.570      1    8    16    0.61319
## phase:hour       1    0.550      0    8     3    0.83245
## treatment:phase:hour  2    0.664      0   16     8    0.99144
## gender:phase:hour  1    0.695      1    8     3    0.62021
## treatment:gender:phase:hour  2    0.793      0   16     8    0.97237
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

# ANCOVA: adding a covariate (necessary to set factorize = FALSE)
aov.car(value ~ treatment * gender + age + Error(id/phase*hour),
        data = obk.long, observed = c("gender", "age"), factorize = FALSE)

aov4(value ~ treatment * gender + age + (phase*hour|id),
      data = obk.long, observed = c("gender", "age"), factorize = FALSE)

ez.glm("id", "value", obk.long, between = c("treatment", "gender"),
       within = c("phase", "hour"), covariate = "age",
       observed = c("gender", "age"), factorize = FALSE)

# aggregating over one within-subjects factor (phase) with warning:
aov.car(value ~ treatment * gender + Error(id/hour), data = obk.long, observed = "gender")

ez.glm("id", "value", obk.long, c("treatment", "gender"), "hour", observed = "gender")

# runs with "numeric" factors
obk.long$hour2 <- as.numeric(as.character(obk.long$hour))

aov.car(value ~ treatment * gender + Error(id/hour2),
        data = obk.long, type = 2, observed = c("gender"))

# only between
aov.car(value ~ treatment * gender + Error(id),
        data = obk.long, observed = c("gender"))
aov4(value ~ treatment * gender + (1|id),
      data = obk.long, observed = c("gender"))
ez.glm("id", "value", obk.long, c("treatment", "gender"),
       within = NULL, print.formula = TRUE, observed = "gender")

# only within
aov.car(value ~ Error(id/phase*hour), data = obk.long, type = 2)

aov4(value ~ (phase*hour|id), data = obk.long, type = 2,
      print.formula = TRUE)

```



```

ez.glm("id", "value", obk.long, NULL, c("phase", "hour"),
      type = 2, print.formula = TRUE)

# using return = "full":

str(aov.car(value ~ Error(id/phase*hour), data = obk.long, return = "full"), 1)

## List of 4
## $ Anova:List of 14
##   ..- attr(*, "class")= chr "Anova.mlm"
## $ lm :List of 11
##   ..- attr(*, "class")= chr [1:2] "mlm" "lm"
## $ data :'data.frame': 16 obs. of 16 variables:
## $ idata:'data.frame': 15 obs. of 2 variables:
## $ marginal:List of 3

# use args.return arguments:
aov.car(value ~ treatment * gender + Error(id/phase*hour),
      data = obk.long, args.return = list(correction = "none", es = "pes"))

aov.car(value ~ treatment * gender + Error(id/phase*hour),
      data = obk.long, observed = "gender",
      args.return = list(correction = "none", MSE = FALSE))

```

compare.2.vectors *Compare two vectors using various tests.*

Description

Compares two vectors x and y using t-test, Welch-test (also known as Satterthwaite), Wilcoxon-test, and a permutation test implemented in **coin**.

Usage

```

compare.2.vectors(x, y, paired = FALSE, na.rm = FALSE,
  tests = c("parametric", "nonparametric"), coin = TRUE,
  alternative = "two.sided",
  perm.distribution = approximate(100000),
  wilcox.exact = NULL, wilcox.correct = TRUE)

```

Arguments

<code>x</code>	a (non-empty) numeric vector of data values.
<code>y</code>	a (non-empty) numeric vector of data values.
<code>paired</code>	a logical whether the data is paired. Default is FALSE.
<code>na.rm</code>	logical. Should NA be removed? Default is FALSE.

tests	Which tests to report, parametric or nonparametric? The default <code>c("parametric", "nonparametric")</code> reports both. See details. (Arguments may be abbreviated).
alternative	a character, the alternative hypothesis must be one of <code>"two.sided"</code> (default), <code>"greater"</code> or <code>"less"</code> . You can specify just the initial letter, will be passed to all functions.
coin	logical or character. Should (permutation) tests from the coin package be reported? Default is <code>TRUE</code> corresponding to all implemented tests. <code>FALSE</code> calculates no tests from coin . A character vector may include any of the following (potentially abbreviated) implemented tests (see also Details): <code>c("permutation", "Wilcoxon", "median")</code> .
perm.distribution	distribution argument to coin , see Distribution or IndependenceTest . Defaults to <code>approximate(100000)</code> indicating an approximation of the exact conditional distribution with 100.000 Monte Carlo samples. One can use <code>"exact"</code> for small samples and if <code>paired = FALSE</code> .
wilcox.exact	exact argument to wilcox.test .
wilcox.correct	correct argument to wilcox.test .

Details

The parametric tests (currently) only contain the *t*-test and Welch/Statterwaithe/Smith/unequal variance *t*-test implemented in [t.test](#). The latter one is only displayed if `paired = FALSE`.

The nonparametric tests (currently) contain the Wilcoxon test implemented in [wilcox.test](#) (`stats::Wilcoxon`) and (if `coin = TRUE`) the following tests implemented in **coin**:

- a permutation test [oneway_test](#) (the only test in this selection not using a rank transformation),
- the Wilcoxon test [wilcox_test](#) (`coin::Wilcoxon`), and
- the median test [median_test](#).

Note that the two implementations of the Wilcoxon test probably differ. This is due to differences in the calculation of the Null distributions.

Value

a list with up to two elements (i.e., parametric and/or nonparametric) each containing a `data.frame` with the following columns: `test`, `test.statistic`, `test.value`, `test.df`, `p`.

Examples

```
with(sleep, compare.2.vectors(extra[group == 1], extra[group == 2]))

# gives:
## $parametric
##   test test.statistic test.value test.df      p
## 1    t              t    -1.861   18.00 0.07919
## 2 Welch              t    -1.861   17.78 0.07939
##
```

```
## $nonparametric
##          test test.statistic test.value test.df      p
## 1 stats::Wilcoxon           W      25.500     NA 0.06933
## 2  permutation             Z      -1.751     NA 0.08154
## 3 coin::Wilcoxon           Z      -1.854     NA 0.06487
## 4      median              Z       1.744     NA 0.17867

# compare with:
with(sleep, compare.2.vectors(extra[group == 1], extra[group == 2], alternative = "less"))

with(sleep, compare.2.vectors(extra[group == 1], extra[group == 2], alternative = "greater"))

# doesn't make much sense as the data is not paired, but whatever:
with(sleep, compare.2.vectors(extra[group == 1], extra[group == 2], paired = TRUE))

# from ?t.test:
compare.2.vectors(1:10,y=c(7:20, 200))
```

md_12.1

Data 12.1 from Maxwell & Delaney

Description

Hypothetical Reaction Time Data for 2 x 3 Perceptual Experiment: Example data for chapter 12 of Maxwell and Delaney (2004, Table 12.1, p. 574) in long format. Has two within.subjects factors: angle and noise.

Usage

```
md_12.1
```

Format

A data.frame with 60 rows and 4 variables.

Details

Description from pp. 573:

Suppose that a perceptual psychologist studying the visual system was interested in determining the extent to which interfering visual stimuli slow the ability to recognize letters. Subjects are brought into a laboratory and seated in front of a tachistoscope. Subjects are told that they will see either the letter T or the letter I displayed on the screen. In some trials, the letter appears by itself, but in other trials, the target letter is embedded in a group of other letters. This variation in the display constitutes the first factor, which is referred to as noise. The noise factor has two levels?absent and present. The other factor varied by the experimenter is where in the display the target letter appears. This factor, which is called angle, has three levels. The target letter is either shown at the center of the screen (i.e., 0° off-center, where the subject has been instructed to fixate), 4° off-center or 8° off-center (in each case, the deviation from the center varies randomly between left and right). Table

12.1 presents hypothetical data for 10 subjects. As usual, the sample size is kept small to make the calculations easier to follow. The dependent measure is reaction time (latency), measured in milliseconds (ms), required by a subject to identify the correct target letter. Notice that each subject has six scores, one for each combination of the 2 x 3 design. In an actual perceptual experiment, each of these six scores would itself be the mean score for that subject across a number of trials in the particular condition. Although "trials" could be used as a third within-subjects factor in such a situation, more typically trials are simply averaged over to obtain a more stable measure of the individual's performance in each condition.

Source

Maxwell, S. E., & Delaney, H. D. (2004). Designing experiments and analyzing data: a model-comparisons perspective. Mahwah, N.J.: Lawrence Erlbaum Associates. p. 574

Examples

```
data(md_12.1)

# Table 12.5 (p. 578):
ez.glm("id", "rt", md_12.1, within = c("angle", "noise"),
      args.return=list(correction = "none", es = "none"))
```

md_15.1

Data 15.1 / 11.5 from Maxwell & Delaney

Description

Hypothetical IQ Data from 12 children at 4 time points: Example data for chapter 11/15 of Maxwell and Delaney (2004, Table 15.1, p. 766) in long format. Has two one within-subjects factor: time.

Usage

md_15.1

Format

A data.frame with 48 rows and 4 variables.

Details

Description from pp. 534:

The data show that 12 subjects have been observed in each of 4 conditions. To make the example easier to discuss, let's suppose that the 12 subjects are children who have been observed at 30, 36, 42, and 48 months of age. In each case, the dependent variable is the child's age-normed general cognitive score on the McCarthy Scales of Children's Abilities. Although the test is normed so that the mean score is independent of age for the general population, our 12 children may come from a population in which cognitive abilities are either growing more rapidly or less rapidly than average. Indeed, this is the hypothesis our data allow us to address. In other words, although the sample

means suggest that the children's cognitive abilities are growing, a significance test is needed if we want to rule out sampling error as a likely explanation for the observed differences.

To replicate the results in chapter 15 several different contrasts need to be applied, see Examples.

time is time in months (centered at 0) and timecat is the same as a categorical variable.

Author(s)

R code for examples written by Ulf Mertens and Henrik Singmann

Source

Maxwell, S. E., & Delaney, H. D. (2004). Designing experiments and analyzing data: a model-comparisons perspective. Mahwah, N.J.: Lawrence Erlbaum Associates. p. 766

Examples

```
### replicate results from Table 15.2 to 15.6 (Maxwell & Delaney, 2004, pp. 774)
data(md_15.1)

### ANOVA results (Table 15.2)
aov4(iq ~ timecat + (timecat|id),data=md_15.1, args.return=list(correction = "none"))

### Table 15.3 (random intercept only)
# we need to set the base level on the last level:
contrasts(md_15.1$timecat) <- contr.treatment(4, base = 4)
# "Type 3 Tests of Fixed Effects"
(t15.3 <- mixed(iq ~ timecat + (1|id),data=md_15.1, check.contrasts=FALSE))
# "Solution for Fixed Effects" and "Covariance Parameter Estimates"
summary(t15.3$full.model)

### make Figure 15.2
plot(NULL, NULL, ylim = c(80, 140), xlim = c(30, 48), ylab = "iq", xlab = "time")
plyr::d_ply(md_15.1, plyr::.id, function(x) lines(as.numeric(as.character(x$timecat)), x$iq))

### Table 15.4, page 789
# random intercept plus slope
(t15.4 <- mixed(iq ~ timecat + (1+time|id),data=md_15.1, check.contrasts=FALSE))
summary(t15.4$full.model)

### Table 15.5, page 795
# set up polynomial contrasts for timecat
contrasts(md_15.1$timecat) <- contr.poly
# fit all parameters separately
(t15.5 <- mixed(iq ~ timecat + (1+time|id), data=md_15.1, check.contrasts=FALSE,
               per.parameter="timecat"))
# quadratic trend is considerably off, conclusions stay the same.

### Table 15.6, page 797
# growth curve model
(t15.6 <- mixed(iq ~ time + (1+time|id),data=md_15.1))
summary(t15.6$full.model)
```

md_16.1

*Data 16.1 / 10.9 from Maxwell & Delaney***Description**

Hypothetical Reaction Time Data for 2 x 3 Perceptual Experiment: Example data for chapter 12 of Maxwell and Delaney (2004, Table 12.1, p. 574) in long format. Has two within.subjects factors: angle and noise.

Usage

md_16.1

Format

A data.frame with 24 rows and 3 variables.

Details

Description from pp. 829:

As brief background, the goal of the study here is to examine the extent to which female and male clinical psychology graduate student trainees may assign different severity ratings to clients at initial intake. Three female and 3 male graduate students are randomly selected to participate and each is randomly assigned four clients with whom to do an intake interview, after which each clinical trainee assigns a severity rating to each client, producing the data shown in Table 16.1.

Note that I changed the labeling of the id slightly, so that they are now labeled from 1 to 6. Furthermore, I changed the contrasts of sex to `contr.treatment` to replicate the exact results of Table 16.3 (p. 837).

Source

Maxwell, S. E., & Delaney, H. D. (2004). *Designing experiments and analyzing data: a model-comparisons perspective*. Mahwah, N.J.: Lawrence Erlbaum Associates. p. 574

Examples

```
### replicate results from Table 16.3 (Maxwell & Delaney, 2004, p. 837)
data(md_16.1)

# original results need treatment contrasts:
(mixed1_orig <- mixed(severity ~ sex + (1|id), md_16.1, check.contrasts=FALSE))
summary(mixed1_orig$full.model)

# p-values stay the same with afex default contrasts (contr.sum),
# but estimates and t-values for the fixed effects parameters change.
(mixed1 <- mixed(severity ~ sex + (1|id), md_16.1))
summary(mixed1$full.model)
```

md_16.4

Data 16.4 from Maxwell & Delaney

Description

Data from a hypothetical inductive reasoning study.

Usage

md_16.4

Format

A data.frame with 24 rows and 3 variables.

Details

Description from pp. 841:

Suppose an educational psychologist has developed an intervention to teach inductive reasoning skills to school children. She decides to test the efficacy of her intervention by conducting a randomized design. Three classrooms of students are randomly assigned to the treatment condition, and 3 other classrooms are assigned to the control.

Table 16.4 shows hypothetical data collected from 29 children who participated in the study assessing the effectiveness of the intervention to increase inductive reasoning skills. We want to call your attention to several aspects of the data. First, the 15 children with condition values of 0 received the control, whereas the 14 children with condition values of 1 received the treatment. Second, 4 of the children in the control condition were students in control Classroom 1, 6 of them were students in control Classroom 2, and 5 were students in control Classroom 3. Along similar lines, 3 of the children in the treatment condition were students in treatment Classroom 1, 5 were students in treatment Classroom 2, and 6 were students in treatment Classroom 3. It is essential to understand that there are a total of six classrooms here; we have coded classroom from 1 to 3 for control as well as treatment, because we will indicate to PROC MIXED that classroom is nested under treatment. Third, scores on the dependent variable appear in the rightmost column under the variable label "induct."

Note that it would make a lot more sense to change the labeling of room from 1 to 3 nested within cond to 1 to 6. However, I keep this in line with the original. The random effects term in the call to mixed is therefore a little bit uncommon.#'

Source

Maxwell, S. E., & Delaney, H. D. (2004). Designing experiments and analyzing data: a model-comparisons perspective. Mahwah, N.J.: Lawrence Erlbaum Associates. p. 574

Examples

```
# data for next examples (Maxwell & Delaney, Table 16.4)
data(md_16.4)
str(md_16.4)

### replicate results from Table 16.6 (Maxwell & Delaney, 2004, p. 845)
# p-values (almost) hold:
(mixed2 <- mixed(induct ~ cond + (1|room:cond), md_16.4))
# (1|room:cond) is needed because room is nested within cond.
```

mixed

Obtain p-values for a mixed-model from lmer().

Description

Fits and calculates p-values for all effects in a mixed model fitted with `lmer`. The default behavior calculates type 3 like p-values using the Kenward-Roger approximation for degrees-of-freedom implemented in `KRmodcomp` (for LMMs only), but also allows for parametric bootstrap (`method = "PB"`), or likelihood ratio tests (the latter two for LMMs and GLMMs). `print`, `summary`, and `anova` methods for the returned object of class "mixed" are available (the last two return the same data.frame).

Usage

```
mixed(formula, data, type = 3, method = c("KR", "PB", "LRT"),
      per.parameter = NULL, args.test = list(), test.intercept = FALSE,
      check.contrasts = TRUE, progress = TRUE, cl = NULL, ...)
```

Arguments

<code>formula</code>	a formula describing the full mixed-model to be fitted. As this formula is passed to <code>lmer</code> , it needs at least one random term.
<code>data</code>	data.frame containing the data. Should have all the variables present in <code>fixed</code> , <code>random</code> , and <code>dv</code> as columns.
<code>type</code>	type of test on which effects are based. Only type 3 tests (3 or "III") are correctly implemented (see Details).
<code>method</code>	character vector indicating which methods for obtaining p-values should be used. "KR" (the default) corresponds to the Kenward-Roger approximation for degrees of freedom (only working with linear mixed models). "PB" calculates p-values based on parametric bootstrap. "LRT" calculates p-values via the likelihood ratio tests implemented in the <code>anova</code> method for <code>merMod</code> objects (only recommended for models with many [i.e., > 50] levels for the random factors).
<code>per.parameter</code>	character vector specifying for which variable tests should be run for each parameter (instead for the overall effect). Can be useful e.g., for testing ordered factors. Relatively untested so results should be compared with a second run without setting this argument. Uses <code>grep</code> for selecting parameters among the fixed effects so regular expressions (<code>regex</code>) are possible. See Examples.

<code>args.test</code>	list of arguments passed to the function calculating the p-values. See Details.
<code>test.intercept</code>	logical. Whether or not the intercept should also be fitted and tested for significance. Default is FALSE. Only relevant if <code>type = 3</code> .
<code>check.contrasts</code>	logical. Should contrasts be checked and (if necessary) changed to "contr.sum"? See Details.
<code>progress</code>	if TRUE, shows progress with a text progress bar and other status messages during fitting.
<code>cl</code>	A vector identifying a cluster; used for distributing the estimation of the different models using several cores. See examples. If <code>check.contrasts</code> , <code>mixed</code> sets the current contrasts (<code>getOption("contrasts")</code>) at the nodes. Note this does <i>not</i> distribute calculation of p-values (e.g., when using <code>method = "PB"</code>) across the cluster. Use <code>args.test</code> for this.
<code>...</code>	further arguments (such as <code>weights</code>) passed to <code>lmer</code> .

Details

For an introduction to mixed-modeling for experimental designs see Barr, Levy, Scheepers, & Tily (2013; I highly recommend reading this paper if you use this function), arguments for using the Kenward-Roger approximation for obtaining p-values are given by Judd, Westfall, and Kenny (2012). Further introductions to mixed-modeling for experimental designs are given by Baayen and colleagues (Baayen, 2008; Baayen, Davidson & Bates, 2008; Baayen & Milin, 2010). Specific recommendations on which random effects structure to specify for confirmatory tests can be found in Barr and colleagues (2013).

p-values are per default calculated via methods from `pbkrtest`. When `method = "KR"` (the default), the Kenward-Roger approximation for degrees-of-freedom is calculated using `KRmodcomp`, which is only applicable to linear-mixed models. The test statistic in the output is a F-value (F).

`method = "PB"` calculates p-values using parametric bootstrap using `PBmodcomp`. This can be used for linear and also generalized linear mixed models (GLMM) by specifying a `family` argument to `mixed`. Note that you should specify further arguments to `PBmodcomp` via `args.test`, especially `nsim` (the number of simulations to form the reference distribution) or `cl` (for using multiple cores). For other arguments see `PBmodcomp`. Note that REML (argument to `[g]lmer`) will be set to FALSE if `method` is PB.

`method = "LRT"` calculates p-values via likelihood ratio tests implemented in the `anova` method for "merMod" objects. This is recommended by Barr et al. (2013; which did not test the other methods implemented here). Using likelihood ratio tests is only recommended for models with many levels for the random effects (> 50), but can be pretty helpful in case the other methods fail (due to memory and/or time limitations). The [lme4 faq](#) also recommends the other methods over likelihood ratio tests.

Type 3 tests are obtained by comparing a model in which only the tested effect is excluded with the full model (containing all effects). This corresponds to the (type 3) Wald tests given by `car::Anova` for "lmerMod" models. The submodels in which the tested effect is excluded are obtained by manually creating a model matrix which is then fitted in "lme4". This is done to avoid R's "feature" to not allow this behavior.

Type 2 tests are truly sequential. They are obtained by comparing a model in which the tested effect and all higher order effect (e.g., all three-way interactions for testing a two-way interaction)

are excluded with a model in which only effects up to the order of the tested effect are present and all higher order effects absent. In other words, there are multiple full models, one for each order of effects. Consequently, the results for lower order effects are identical of whether or not higher order effects are part of the model or not. This latter feature is not consistent with classical ANOVA type 2 tests but a consequence of the sequential tests (and **I didn't find a better way** of implementing the Type 2 tests). This **does not** correspond to the (type 2) Wald test reported by `car::Anova`. If you want type 2 Wald tests instead of truly sequential type 2 tests, use `car::Anova` with `test = "F"`. Note that the order in which the effects are entered into the formula does not matter (in contrast to type 1 tests).

If `check.contrasts = TRUE`, contrasts will be set to `"contr.sum"` for all factors in the formula if default contrasts are not equal to `"contr.sum"` or `attrib(factor, "contrasts") != "contr.sum"`. Furthermore, the current contrasts (obtained via `getOption("contrasts")`) will be set at the cluster nodes if `cl` is not `NULL`.

Value

An object of class `"mixed"` (i.e., a list) with the following elements:

1. `anova.table` a data.frame containing the statistics returned from `KRmodcomp`. The `stat` column in this data.frame gives the value of the test statistic, an F-value for `method = "KR"` and a chi-square value for the other two methods.
2. `full.model` the `"lmerMod"` object returned from fitting the full mixed model.
3. `restricted.models` a list of `"lmerMod"` objects from fitting the restricted models (i.e., each model lacks the corresponding effect)
4. `tests` a list of objects returned by the function for obtaining the p-values.
5. `type` The `type` argument used when calling this function.
6. `method` The `method` argument used when calling this function.

The following methods exist for objects of class `"mixed"`: `print` (which uses rounding and invisibly returns the output), `summary`, and `anova` (the latter two return the same data.frame).

Note

When `method = "KR"`, obtaining p-values is known to crash due to insufficient memory or other computational limitations (especially with complex random effects structures). In these cases, the other methods should be used. The RAM demand is a problem especially on 32 bit Windows which only supports up to 2 or 3GB RAM (see [R Windows FAQ](#)). Then it is probably a good idea to use methods `"LRT"` or `"PB"`.

`"mixed"` will throw a message if numerical variables are not centered on 0, as main effects (of other variables than the numeric one) can be hard to interpret if numerical variables appear in interactions. See

Please report all bugs to `henrik.singmann (at) psychologie.uni-freiburg.de`

Author(s)

Henrik Singmann with contributions from [Ben Bolker and Joshua Wiley](#).

References

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- Judd, C. M., Westfall, J., & Kenny, D. A. (2012). Treating stimuli as a random factor in social psychology: A new and comprehensive solution to a pervasive but largely ignored problem. *Journal of Personality and Social Psychology*, 103(1), 54-69. doi:10.1037/a0028347
- Maxwell, S. E., & Delaney, H. D. (2004). *Designing experiments and analyzing data: a model-comparisons perspective*. Mahwah, N.J.: Lawrence Erlbaum Associates.

See Also

[ez.glm](#) and [aov.car](#) for convenience functions to analyze experimental designs with classical ANOVA or ANCOVA wrapping [Anova](#).

see the following for the data sets from Maxwell and Delaney (2004) used and more examples: [md_15.1](#), [md_16.1](#), and [md_16.4](#).

Examples

```
### replicate results from Table 15.4 (Maxwell & Delaney, 2004, p. 789)
data(md_15.1)
# random intercept plus slope
(t15.4a <- mixed(iq ~ timecat + (1+time|id),data=md_15.1))

# to also replicate exact parameters use treatment.contrasts and the last level as base level:
contrasts(md_15.1$timecat) <- contr.treatment(4, base = 4)
(t15.4b <- mixed(iq ~ timecat + (1+time|id),data=md_15.1, check.contrasts=FALSE))
summary(t15.4a$full.model) # gives "wrong" parameters estimates
summary(t15.4b$full.model) # identical parameters estimates

# for more examples from chapter 15 see ?md_15.1

### replicate results from Table 16.3 (Maxwell & Delaney, 2004, p. 837)
data(md_16.1)

# original results need treatment contrasts:
(mixed1_orig <- mixed(severity ~ sex + (1|id), md_16.1, check.contrasts=FALSE))
summary(mixed1_orig$full.model)
```

```

# p-value stays the same with afex default contrasts (contr.sum),
# but estimates and t-values for the fixed effects parameters change.
(mixed1 <- mixed(severity ~ sex + (1|id), md_16.1))
summary(mixed1$full.model)

# data for next examples (Maxwell & Delaney, Table 16.4)
data(md_16.4)
str(md_16.4)

### replicate results from Table 16.6 (Maxwell & Delaney, 2004, p. 845)
# Note that (1|room:cond) is needed because room is nested within cond.
# p-value (almost) holds.
(mixed2 <- mixed(induct ~ cond + (1|room:cond), md_16.4))
# (differences are due to the use of Kenward-Roger approximation here,
# whereas M&W's p-values are based on uncorrected df.)

# again, to obtain identical parameter and t-values, use treatment contrasts:
summary(mixed2$full.model) # not identical

# prepare new data.frame with contrasts:
md_16.4b <- within(md_16.4, cond <- C(cond, contr.treatment, base = 2))
str(md_16.4b)

# p-value stays identical:
(mixed2_orig <- mixed(induct ~ cond + (1|room:cond), md_16.4b, check.contrasts=FALSE))
summary(mixed2_orig$full.model) # replicates parameters

### replicate results from Table 16.7 (Maxwell & Delaney, 2004, p. 851)
# F-values (almost) hold, p-values (especially for skill) are off
(mixed3 <- mixed(induct ~ cond + skill + (1|room:cond), md_16.4))

# however, parameters are perfectly recovered when using the original contrasts:
mixed3_orig <- mixed(induct ~ cond + skill + (1|room:cond), md_16.4b, check.contrasts=FALSE)
summary(mixed3_orig$full.model)

### replicate results from Table 16.10 (Maxwell & Delaney, 2004, p. 862)
# for this we need to center cog:
md_16.4b$cog <- scale(md_16.4b$cog, scale=FALSE)

# F-values and p-values are relatively off:
(mixed4 <- mixed(induct ~ cond*cog + (cog|room:cond), md_16.4b))
# contrast has a relatively important influence on cog
(mixed4_orig <- mixed(induct ~ cond*cog + (cog|room:cond), md_16.4b, check.contrasts=FALSE))

# parameters are again almost perfectly recovered:
summary(mixed4_orig$full.model)

```

```

## Not run:

# use the obk.long data (not reasonable, no random slopes)
data(obk.long)
mixed(value ~ treatment * phase + (1|id), obk.long)

# Examples for using the per.parameter argument:
data(obk.long, package = "afex")
obk.long$hour <- ordered(obk.long$hour)

# tests only the main effect parameters of hour individually per parameter.
mixed(value ~ treatment*phase*hour +(1|id), per.parameter = "^hour$", data = obk.long)

# tests all parameters including hour individually
mixed(value ~ treatment*phase*hour +(1|id), per.parameter = "hour", data = obk.long)

# tests all parameters individually
mixed(value ~ treatment*phase*hour +(1|id), per.parameter = ".", data = obk.long)

# example data from package languageR:
# Lexical decision latencies elicited from 21 subjects for 79 English concrete nouns,
# with variables linked to subject or word.
data(lexdec, package = "languageR")

# using the simplest model
m1 <- mixed(RT ~ Correct + Trial + PrevType * meanWeight +
  Frequency + NativeLanguage * Length + (1|Subject) + (1|Word), data = lexdec)

m1
##           Effect  stat ndf    ddf F.scaling p.value
## 1           Correct  8.15  1 1627.73     1.00   .004
## 2             Trial  7.57  1 1592.43     1.00   .006
## 3           PrevType  0.17  1 1605.39     1.00   .68
## 4         meanWeight 14.85  1   75.39     1.00  .0002
## 5           Frequency 56.53  1   76.08     1.00 <.0001
## 6   NativeLanguage  0.70  1   27.11     1.00   .41
## 7             Length  8.70  1   75.83     1.00   .004
## 8 PrevType:meanWeight  6.18  1 1601.18     1.00   .01
## 9 NativeLanguage:Length 14.24  1 1555.49     1.00  .0002

# Fitting a GLMM using parametric bootstrap:
require("mlmRev") # for the data, see ?Contraception

gm1 <- mixed(use ~ age + I(age^2) + urban + livch + (1 | district), method = "PB",
  family = binomial, data = Contraception, args.test = list(nsim = 10))

#####
### using multicore ###
#####

require(parallel)
(nc <- detectCores()) # number of cores
cl <- makeCluster(rep("localhost", nc)) # make cluster

```

```

# to keep track of what the function is doing redirect output to outfile:
# cl <- makeCluster(rep("localhost", nc), outfile = "cl.log.txt")

## There are two ways to use multicore:

# 1. Obtain fits with multicore:
mixed(value ~ treatment*phase*hour +(1|id), data = obk.long, method = "LRT", cl = cl)

# 2. Obtain PB samples via multicore:
mixed(use ~ age + I(age^2) + urban + livch + (1 | district), family = binomial,
      method = "PB", data = Contraception, args.test = list(nsim = 10, cl = cl))

## Both ways can be combined:
mixed(use ~ age + I(age^2) + urban + livch + (1 | district), family = binomial,
      method = "PB", data = Contraception, args.test = list(nsim = 10, cl = cl), cl = cl)

stopCluster(cl)

## End(Not run)

```

nice.anova

Make nice ANOVA table for printing.

Description

These functions produce a nice ANOVA table best for printing. `nice.anova` takes an object from [Anova](#) possible created by the convenience functions [ez.glm](#) or [aov.car](#). When within-subject factors are present, either sphericity corrected or uncorrected degrees of freedom can be reported.

Usage

```
nice.anova(object, es = "ges", observed = NULL, correction = c("GG", "HF", "none"),
           MSE = TRUE, intercept = FALSE, sig.symbols = c("+", " *", " **", " ***"))
```

Arguments

object	An object of class "Anova.mlm" or "anova" as returned from Anova , ez.glm , or aov.car .
es	Effect Size to be reported. Default is "ges", which reports generalized eta-squared (see details). Also supported is partial eta-squared ("pes") or "none".
observed	character vector referring to the observed (i.e., non manipulated) variables/effects in the design. Important for calculation of generalized eta-squared (ignored if es is not "ges"), see details.
correction	Character. Which sphericity correction on the degrees of freedom should be reported for the within-subject factors. The default <code>c("GG", "HF", "none")</code> corresponds to the Greenhouse-Geisser correction.

sig.symbols	Character. What should be the symbols designating significance? When entering an vector with <code>length(sig.symbol) < 4</code> only those elements of the default (<code>c(" +", " *", " **", " ***")</code>) will be replaced. <code>sig.symbols = ""</code> will display the stars but not the +, <code>sig.symbols = rep("", 4)</code> will display no symbols.
MSE	logical. Should the column containing the Mean Squared Error (MSE) be displayed? Default is TRUE.
intercept	logical. Should intercept (if present) be printed (default is FALSE which suppresses printing of the intercept)

Details

The returned data.frame is print-ready when adding to a document with proper methods. I recommend `ascii` and `xtable`. `ascii` provides conversion to `AsciiDoc` but most notably to `org-mode` (see `ascii` and `print-ascii`). `xtable` converts a data.frame into LaTeX code with many possible options (e.g., allowing for "longtable" or "sidewaystable"), see `xtable` and `print.xtable`. See Examples.

Conversion functions to other formats (such as HTML, ODF, or Word) can be found at the [Reproducible Research Task View](#).

The default reports generalized eta squared (Olejnik & Algina, 2003), the "recommended effect size for repeated measured designs" (Bakeman, 2005). Note that it is important that all measured variables (as opposed to experimentally manipulated variables), such as e.g., age, gender, weight, ..., must be declared via `observed` to obtain the correct effect size estimate. Partial eta squared ("pes") does not require this.

Value

A data.frame with the ANOVA table consisting of characters. The columns that are always present are: Effect, df (degrees of freedom), F, and p.

`ges` contains the generalized eta-squared effect size measure (Bakeman, 2005), `pes` contains partial eta-squared (if requested).

Author(s)

The code for calculating generalized eta-squared was written by Mike Lawrence. Everything else was written by Henrik Singmann.

References

- Bakeman, R. (2005). Recommended effect size statistics for repeated measures designs. *Behavior Research Methods*, 37(3), 379-384. doi:10.3758/BF03192707
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See Also

`ez.glm` and `aov.car` are the convenience functions to create the object appropriate for `nice.anova`.

Examples

```
## example from Olejnik & Algina (2003)
# "Repeated Measures Design" (pp. 439):
data(md_12.1)
# create object of class Anova:
rmd <- ez.glm("id", "rt", md_12.1, within = c("angle", "noise"),
             return = "Anova")
# use different es:
nice.anova(rmd, es = "pes") # noise: .82
nice.anova(rmd, es = "ges") # noise: .39

# exampel using obk.long (see ?obk.long), a long version of the OBrienKaiser dataset from car.
data(obk.long)
# create object of class Anova:
tmp.aov <- aov.car(value ~ treatment * gender + Error(id/phase*hour),
                  data = obk.long, return = "Anova")

nice.anova(tmp.aov, observed = "gender")

nice.anova(tmp.aov, observed = "gender", sig.symbol = rep("", 4))

## Not run:
# use package ascii or xtable for formatting of tables ready for printing.

full <- nice.anova(tmp.aov, observed = "gender")

require(ascii)
print(ascii(full, include.rownames = FALSE, caption = "ANOVA 1"), type = "org")

require(xtable)
print.xtable(xtable(full, caption = "ANOVA 2"), include.rownames = FALSE)

## End(Not run)
```

obk.long

O'Brien Kaiser's Repeated-Measures Dataset with Covariate

Description

This is the long version of the OBrienKaiser dataset from the **car** package adding a random covariate age. Originally the dataset ist taken from O'Brien and Kaiser (1985). The description from [OBrienKaiser](#) says: "These contrived repeated-measures data are taken from O'Brien and Kaiser (1985). The data are from an imaginary study in which 16 female and male subjects, who are divided into three treatments, are measured at a pretest, posttest, and a follow-up session; during each session, they are measured at five occasions at intervals of one hour. The design, therefore, has two between-subject and two within-subject factors."

Usage

obk.long

Format

A data frame with 240 rows and 7 variables.

Source

O'Brien, R. G., & Kaiser, M. K. (1985). MANOVA method for analyzing repeated measures designs: An extensive primer. *Psychological Bulletin*, 97, 316-333. doi:10.1037/0033-2909.97.2.316

Examples

```
# The dataset is constructed as follows:
set.seed(1)
OBrienKaiser2 <- within(OBrienKaiser, {
  id <- factor(1:nrow(OBrienKaiser))
  age <- scale(sample(18:35, nrow(OBrienKaiser), replace = TRUE), scale = FALSE))
  attributes(OBrienKaiser2$age) <- NULL # needed or reshape2::melt throws an error.
  OBrienKaiser2$age <- as.numeric(OBrienKaiser2$age)
  obk.long <- reshape2::melt(OBrienKaiser2, id.vars = c("id", "treatment", "gender", "age"))
  obk.long[,c("phase", "hour")] <- lapply(as.data.frame(do.call(rbind,
    strsplit(as.character(obk.long$variable), "\\."),)), factor)
  obk.long <- obk.long[,c("id", "treatment", "gender", "age", "phase", "hour", "value")]
  obk.long <- obk.long[order(obk.long$id),]
  rownames(obk.long) <- NULL
  str(obk.long)
  ## 'data.frame': 240 obs. of 7 variables:
  ## $ id : Factor w/ 16 levels "1","2","3","4",...: 1 1 1 1 1 1 1 1 1 1 ...
  ## $ treatment: Factor w/ 3 levels "control","A",...: 1 1 1 1 1 1 1 1 1 1 ...
  ## $ gender : Factor w/ 2 levels "F","M": 2 2 2 2 2 2 2 2 2 2 ...
  ## $ age : num -4.75 -4.75 -4.75 -4.75 -4.75 -4.75 -4.75 -4.75 -4.75 -4.75 ...
  ## $ phase : Factor w/ 3 levels "fup","post","pre": 3 3 3 3 3 2 2 2 2 2 ...
  ## $ hour : Factor w/ 5 levels "1","2","3","4",...: 1 2 3 4 5 1 2 3 4 5 ...
  ## $ value : num 1 2 4 2 1 3 2 5 3 2 ...
  head(obk.long)
  ## id treatment gender age phase hour value
  ## 1 1 control M -4.75 pre 1 1
  ## 2 1 control M -4.75 pre 2 2
  ## 3 1 control M -4.75 pre 3 4
  ## 4 1 control M -4.75 pre 4 2
  ## 5 1 control M -4.75 pre 5 1
  ## 6 1 control M -4.75 post 1 3
```

round_ps

Helper function which rounds p-values

Description

p-values are rounded in a sane way: .99 - .01 to two digits, < .01 to three digits, < .001 to four digits.

Usage

```
round_ps(x)
```

Arguments

x a numeric vector

Value

A character vector with the same length of x.

Author(s)

Henrik Singmann

Examples

```
round_ps(runif(10))
```

```
round_ps(runif(10, 0, .01))
```

```
round_ps(runif(10, 0, .001))
```

```
round_ps(0.0000000099)
```

set_sum_contrasts *Set global contrasts*

Description

These functions are simple wrappers to set contrasts globally via `options(contrasts = ...)`.

Usage

```
set_sum_contrasts()
```

```
set_deviation_contrasts()
```

```
set_effects_contrasts()
```

```
set_default_contrasts()
```

```
set_treatment_contrasts()
```

Details

`set_deviation_contrasts` and `set_effects_contrasts` are wrappers for `set_sum_contrasts`. Likewise, `set_default_contrasts` is a wrapper to `set_treatment_contrasts()`.

Value

nothing. These functions are called for their side effects to change the global options.

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