

Package ‘smoothtail’

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Description Given independent and identically distributed observations $X(1), \dots, X(n)$ from a Generalized Pareto distribution with shape parameter γ in $[-1,0]$, offers several estimates to compute estimates of γ . The estimates are based on the principle of replacing the order statistics by quantiles of a distribution function based on a log--concave density function. This procedure is justified by the fact that the GPD density is log--concave for γ in $[-1,0]$.

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smoothtail-package *Smooth Estimation of GPD Shape Parameter*

Description

Given independent and identically distributed observations $X_1 < \dots < X_n$ from a Generalized Pareto distribution with shape parameter $\gamma \in [-1, 0]$, offers three methods to compute estimates of γ . The estimates are based on the principle of replacing the order statistics $X_{(1)}, \dots, X_{(n)}$ of the sample by quantiles $\hat{X}_{(1)}, \dots, \hat{X}_{(n)}$ of the distribution function \hat{F}_n based on the log-concave density estimator \hat{f}_n . This procedure is justified by the fact that the GPD density is log-concave for $\gamma \in [-1, 0]$.

Details

Package: smoothtail
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Use this package to estimate the shape parameter γ of a Generalized Pareto Distribution (GPD). In extreme value theory, γ is denoted tail index. We offer three new estimators, all based on the fact that the density function of the GPD is log-concave if $\gamma \in [-1, 0]$, see Mueller and Rufibach (2009). The functions for estimation of the tail index are:

[pickands](#)
[falk](#)
[falkMVUE](#)
[generalizedPick](#)

This package depends on the package **logcondens** for estimation of a log-concave density: all the above functions take as first argument a `dlc` object as generated by `logConDens` in **logcondens**.

Additionally, functions for density, distribution function, quantile function and random number generation for a GPD with location parameter 0, shape parameter γ and scale parameter σ are provided:

[dgpd](#)
[pgpd](#)
[qgpd](#)
[rgpd](#).

Let us shortly clarify what we mean with log-concave density estimation. Suppose we are given an ordered sample $Y_1 < \dots < Y_n$ of i.i.d. random variables having density function f , where $f = \exp \varphi$ for a concave function $\varphi : [-\infty, \infty) \rightarrow R$. Following the development in Duembgen and Rufibach (2009), it is then possible to get an estimator $\hat{f}_n = \exp \hat{\varphi}_n$ of f via the maximizer $\hat{\varphi}_n$ of

$$L(\varphi) = \sum_{i=1}^n \varphi(Y_i) - \int \exp \varphi(t) dt$$

over all concave functions φ . It turns out that $\hat{\varphi}_n$ is piecewise linear, with knots only at (some of the) observation points. Therefore, the infinite-dimensional optimization problem of finding the function $\hat{\varphi}_n$ boils down to a finite dimensional problem of finding the vector $(\hat{\varphi}_n(Y_1), \dots, \hat{\varphi}_n(Y_n))$. How to solve this problem is described in Rufibach (2006, 2007) and in a more general setting in Duembgen, Huesler, and Rufibach (2010). The distribution function based on \hat{f}_n is defined as

$$\hat{F}_n(x) = \int_{Y_1}^x \hat{f}_n(t) dt$$

for x a real number. The definition of \hat{F}_n is justified by the fact that $\hat{F}_n(Y_1) = 0$.

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Kaspar Rufibach acknowledges support by the Swiss National Science Foundation SNF, <http://www.snf.ch>

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- Mueller, S. and Rufibach K. (2009). Smooth tail index estimation. *J. Stat. Comput. Simul.*, **79**, 1155–1167.
- Mueller, S. and Rufibach K. (2008). On the max-domain of attraction of distributions with log-concave densities. *Statist. Probab. Lett.*, **78**, 1440–1444.
- Rufibach K. (2006) *Log-concave Density Estimation and Bump Hunting for i.i.d. Observations*. PhD Thesis, University of Bern, Switzerland and Georg-August University of Goettingen, Germany, 2006. Available at http://www.zb.unibe.ch/download/eldiss/06rufibach_k.pdf.
- Rufibach, K. (2007) Computing maximum likelihood estimators of a log-concave density function. *J. Stat. Comput. Simul.*, **77**, 561–574.

See Also

Package **logcondens**.

Examples

```
# generate ordered random sample from GPD
set.seed(1977)
n <- 20
gam <- -0.75
x <- rgpd(n, gam)

# compute known endpoint
omega <- -1 / gam

# estimate log-concave density, i.e. generate dlc object
est <- logConDens(x, smoothed = FALSE, print = FALSE, gam = NULL, xs = NULL)

# plot distribution functions
s <- seq(0.01, max(x), by = 0.01)
plot(0, 0, type = 'n', ylim = c(0, 1), xlim = range(c(x, s))); rug(x)
lines(s, pgpd(s, gam), type = 'l', col = 2)
lines(x, 1:n / n, type = 's', col = 3)
lines(x, est$Fhat, type = 'l', col = 4)
legend(1, 0.4, c('true', 'empirical', 'estimated'), col = c(2 : 4), lty = 1)

# compute tail index estimators for all sensible indices k
falk.logcon <- falk(est)
falkMVUE.logcon <- falkMVUE(est, omega)
pick.logcon <- pickands(est)
genPick.logcon <- generalizedPick(est, c = 0.75, gam0 = -1/3)

# plot smoothed and unsmoothed estimators versus number of order statistics
plot(0, 0, type = 'n', xlim = c(0,n), ylim = c(-1, 0.2))
lines(1:n, pick.logcon[, 2], col = 1); lines(1:n, pick.logcon[, 3], col = 1, lty = 2)
lines(1:n, falk.logcon[, 2], col = 2); lines(1:n, falk.logcon[, 3], col = 2, lty = 2)
lines(1:n, falkMVUE.logcon[,2], col = 3); lines(1:n, falkMVUE.logcon[,3], col = 3,
      lty = 2)
lines(1:n, genPick.logcon[, 2], col = 4); lines(1:n, genPick.logcon[, 3], col = 4,
      lty = 2)
abline(h = gam, lty = 3)
legend(11, 0.2, c("Pickands", "Falk", "Falk MVUE", "Generalized Pickands"),
      lty = 1, col = 1:8)
```

falk

Compute original and smoothed version of Falk's estimator

Description

Given an ordered sample of either exceedances or upper order statistics which is to be modeled using a GPD, this function provides Falk's estimator of the shape parameter $\gamma \in [-1, 0]$. Precisely,

$$\hat{\gamma}_{\text{Falk}} = \hat{\gamma}_{\text{Falk}}(k, n) = \frac{1}{k-1} \sum_{j=2}^k \log\left(\frac{X_{(n)} - H^{-1}((n-j+1)/n)}{X_{(n)} - H^{-1}((n-k)/n)}\right), \quad k = 3, \dots, n-1$$

for γ either the empirical or the distribution function based on the log-concave density estimator. Note that for any k , $\hat{\gamma}_{\text{Falk}} : R^n \rightarrow (-\infty, 0)$. If $\hat{\gamma}_{\text{Falk}} \notin [-1, 0)$, then it is likely that the log-concavity assumption is violated.

Usage

```
falk(est, ks = NA)
```

Arguments

<code>est</code>	Log-concave density estimate based on the sample as output by <code>logConDens</code> (a <code>dlc</code> object).
<code>ks</code>	Indices k at which Falk's estimate should be computed. If set to <code>NA</code> defaults to $3, \dots, n - 1$.

Value

$n \times 3$ matrix with columns: indices k , Falk's estimator based on the log-concave density estimate, and the ordinary Falk's estimator based on the order statistics.

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References

- Mueller, S. and Rufibach K. (2009). Smooth tail index estimation. *J. Stat. Comput. Simul.*, **79**, 1155–1167.
- Falk, M. (1995). Some best parameter estimates for distributions with finite endpoint. *Statistics*, **27**, 115–125.

See Also

Other approaches to estimate γ based on the fact that the density is log-concave, thus $\gamma \in [-1, 0]$, are available as the functions `pickands`, `falkMVUE`, `generalizedPick`.

Examples

```
# generate ordered random sample from GPD
set.seed(1977)
n <- 20
gam <- -0.75
x <- rgpd(n, gam)
```

```
## generate dlc object
est <- logConDens(x, smoothed = FALSE, print = FALSE, gam = NULL, xs = NULL)

# compute tail index estimator
falk(est)
```

falkMVUE	<i>Compute original and smoothed version of Falk's estimator for a known endpoint</i>
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Description

Given an ordered sample of either exceedances or upper order statistics which is to be modeled using a GPD with distribution function F , this function provides Falk's estimator of the shape parameter $\gamma \in [-1, 0]$ if the endpoint

$$\omega(F) = \sup\{x : F(x) < 1\}$$

of F is known. Precisely,

$$\hat{\gamma}_{\text{MVUE}} = \hat{\gamma}_{\text{MVUE}}(k, n) = \frac{1}{k} \sum_{j=1}^k \log\left(\frac{\omega(F) - H^{-1}((n-j+1)/n)}{\omega(F) - H^{-1}((n-k)/n)}\right), \quad k = 2, \dots, n-1$$

for H either the empirical or the distribution function based on the log-concave density estimator. Note that for any k , $\hat{\gamma}_{\text{MVUE}} : R^n \rightarrow (-\infty, 0)$. If $\hat{\gamma}_{\text{MVUE}} \notin [-1, 0)$, then it is likely that the log-concavity assumption is violated.

Usage

```
falkMVUE(est, omega, ks = NA)
```

Arguments

est	Log-concave density estimate based on the sample as output by logConDens (a dlc object).
omega	Known endpoint. Make sure that $\omega \geq X_{(n)}$.
ks	Indices k at which Falk's estimate should be computed. If set to NA defaults to $2, \dots, n-1$.

Value

$n \times 3$ matrix with columns: indices k , Falk's MVUE estimator using the log-concave density estimate, and the ordinary Falk MVUE estimator based on the order statistics.

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References

Mueller, S. and Rufibach K. (2009). Smooth tail index estimation. *J. Stat. Comput. Simul.*, **79**, 1155–1167.

Falk, M. (1994). Extreme quantile estimation in δ -neighborhoods of generalized Pareto distributions. *Statistics and Probability Letters*, **20**, 9–21.

Falk, M. (1995). Some best parameter estimates for distributions with finite endpoint. *Statistics*, **27**, 115–125.

See Also

Other approaches to estimate γ based on the fact that the density is log-concave, thus $\gamma \in [-1, 0]$, are available as the functions [pickands](#), [falk](#), [generalizedPick](#).

Examples

```
# generate ordered random sample from GPD
set.seed(1977)
n <- 20
gam <- -0.75
x <- rgpd(n, gam)

## generate dlc object
est <- logConDens(x, smoothed = FALSE, print = FALSE, gam = NULL, xs = NULL)

# compute tail index estimators
omega <- -1 / gam
falkMVUE(est, omega)
```

generalizedPick

Compute generalized Pickand's estimator

Description

Given an ordered sample of either exceedances or upper order statistics which is to be modeled using a GPD with distribution function F , this function provides Segers' estimator of the shape parameter γ , see Segers (2005). Precisely, for $k = \{1, \dots, n - 1\}$, the estimator can be written as

$$\hat{\gamma}_{\text{Segers}}^k(H) = \sum_{j=1}^k \left(\lambda(j/k) - \lambda((j-1)/k) \right) \log \left(H^{-1}((n - \lfloor cj \rfloor)/n) - H^{-1}((n-j)/n) \right)$$

for H either the empirical or the distribution function based on the log-concave density estimator and λ the mixing measure given in Segers (2005), Theorem 4.1, (i). Note that for any k , $\hat{\gamma}_{\text{Segers}}^k : R^n \rightarrow (-\infty, \infty)$. If $\hat{\gamma}_{\text{Segers}} \notin [-1, 0)$, then it is likely that the log-concavity assumption is violated.

Usage

```
generalizedPick(est, c, gam0, ks = NA)
```

Arguments

est	Log-concave density estimate based on the sample as output by logConDens (a dlc object).
c	Number in $(0, 1)$, determining the spacings that are used.
gam0	Number in $R \setminus 0.5$, specifying the mixing measure.
ks	Indices k at which Falk's estimate should be computed. If set to NA defaults to $4, \dots, n$.

Value

$n \times 3$ matrix with columns: indices k , Segers' estimator using the smoothing method, and the ordinary Segers' estimator based on the order statistics.

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References

- Mueller, S. and Rufibach K. (2009). Smooth tail index estimation. *J. Stat. Comput. Simul.*, **79**, 1155–1167.
- Segers, J. (2005). Generalized Pickands estimators for the extreme value index. *J. Statist. Plann. Inference*, **128**, 381–396.

See Also

Other approaches to estimate γ based on the fact that the density is log-concave, thus $\gamma \in [-1, 0]$, are available as the functions [pickands](#), [falk](#), [falkMVUE](#).

Examples

```
# generate ordered random sample from GPD
set.seed(1977)
n <- 20
gam <- -0.75
x <- rgpd(n, gam)

## generate dlc object
est <- logConDens(x, smoothed = FALSE, print = FALSE, gam = NULL, xs = NULL)

# compute tail index estimators
generalizedPick(est, c = 0.75, gam0 = -1/3)
```

gpd

*The Generalized Pareto Distribution***Description**

Density function, distribution function, quantile function and random generation for the generalized Pareto distribution (GPD) with shape parameter γ and scale parameter σ .

Usage

```
dgpd(x, gam, sigma = 1)
pgpd(q, gam, sigma = 1)
qgpd(p, gam, sigma = 1)
rgpd(n, gam, sigma = 1)
```

Arguments

x, q	Vector of quantiles.
p	Vector of probabilities.
n	Number of observations.
gam	Shape parameter, real number.
sigma	Scale parameter, positive real number.

Details

The generalized Pareto distribution function (Pickands, 1975) with shape parameter γ and scale parameter σ is

$$W_{\gamma,\sigma}(x) = 1 - (1 + \gamma x/\sigma)_+^{-1/\gamma}.$$

If $\gamma = 0$, the distribution function is defined by continuity. The density is denoted by $w_{\gamma,\sigma}$.

Value

`dgpd` gives the values of the density function, `pgpd` those of the distribution function, and `qgpd` those of the quantile function of the GPD at x , q , and p , respectively. `rgpd` generates n random numbers, returned as an ordered vector.

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References

Pickands, J. (1975). Statistical inference using extreme order statistics. *Annals of Statistics*, **3**, 119-131.

See Also

Similar functions are provided in the R-packages `evir` and `evd`.

lambdaGenPick

Auxiliary function to compute Segers' estimator

Description

This function computes

$$\lambda_{\delta, \rho}^c$$

given in Theorem 4.1 of Segers (2005) and is called by `generalizedPick`. It is not intended to be called by the user.

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- Segers, J. (2005). Generalized Pickands estimators for the extreme value index. *J. Statist. Plann. Inference*, **128**, 381–396.

See Also

Called by [generalizedPick](#).

pickands

Compute original and smoothed version of Pickands' estimator

Description

Given an ordered sample of either exceedances or upper order statistics which is to be modeled using a GPD, this function provides Pickands' estimator of the shape parameter $\gamma \in [-1, 0]$. Precisely, for $k = 4, \dots, n$

$$\hat{\gamma}_{\text{Pick}}^k = \frac{1}{\log 2} \log \left(\frac{H^{-1}((n - r_k(H) + 1)/n) - H^{-1}((n - 2r_k(H) + 1)/n)}{H^{-1}((n - 2r_k(H) + 1)/n) - H^{-1}((n - 4r_k(H) + 1)/n)} \right)$$

for H either the empirical or the distribution function \hat{F}_n based on the log-concave density estimator and

$$r_k(H) = \lfloor k/4 \rfloor$$

if H is the empirical distribution function and

$$r_k(H) = k/4$$

if $H = \hat{F}_n$.

Usage

`pickands(est, ks = NA)`

Arguments

- `est` Log-concave density estimate based on the sample as output by `logConDens` (a `dlc` object).
- `ks` Indices k at which Falk's estimate should be computed. If set to `NA` defaults to $4, \dots, n$.

Value

$n \times 3$ matrix with columns: indices k , Pickands' estimator using the log-concave density estimate, and the ordinary Pickands' estimator based on the order statistics.

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Pickands, J. (1975). Statistical inference using extreme order statistics. *Annals of Statistics* **3**, 119–131.

See Also

Other approaches to estimate γ based on the fact that the density is log-concave, thus $\gamma \in [-1, 0]$, are available as the functions [falk](#), [falkMVUE](#), [generalizedPick](#).

Examples

```
# generate ordered random sample from GPD
set.seed(1977)
n <- 20
gam <- -0.75
x <- rgpd(n, gam)

## generate dlc object
est <- logConDens(x, smoothed = FALSE, print = FALSE, gam = NULL, xs = NULL)

# compute tail index estimators
pickands(est)
```

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