

Package ‘mcmc’

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Title Markov Chain Monte Carlo

Author Charles J. Geyer <charlie@stat.umn.edu> and Leif T. Johnson
<ltjohnson@google.com>

Maintainer Charles J. Geyer <charlie@stat.umn.edu>

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Suggests xtable, Iso

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Description Simulates continuous distributions of random vectors using Markov chain Monte Carlo (MCMC). Users specify the distribution by an R function that evaluates the log unnormalized density. Algorithms are random walk Metropolis algorithm (function `metrop`), simulated tempering (function `temper`), and morphometric random walk Metropolis (Johnson and Geyer, 2012, <doi:10.1214/12-AOS1048>, function `morph.metrop`), which achieves geometric ergodicity by change of variable.

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URL <http://www.stat.umn.edu/geyer/mcmc/>,
<https://github.com/cjgeyer/mcmc>

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foo	<i>Simulated logistic regression data.</i>
-----	--

Description

Like it says

Usage

```
data(foo)
```

Format

A data frame with variables

x1 quantitative predictor.

x2 quantitative predictor.

x3 quantitative predictor.

y Bernoulli response.

Examples

```
library(mcmc)
data(foo)
out <- glm(y ~ x1 + x2 + x3, family = binomial, data = foo)
summary(out)
```

initseq	<i>Initial Sequence Estimators</i>
---------	------------------------------------

Description

Variance of sample mean of functional of reversible Markov chain using methods of Geyer (1992).

Usage

```
initseq(x)
```

Arguments

`x` a numeric vector that is a scalar-valued functional of a reversible Markov chain.

Details

Let

$$\gamma_k = \text{cov}(X_i, X_{i+k})$$

considered as a function of the lag k be the autocovariance function of the input time series. Define

$$\Gamma_k = \gamma_{2k} + \gamma_{2k+1}$$

the sum of consecutive pairs of autocovariances. Then Theorem 3.1 in Geyer (1992) says that Γ_k considered as a function of k is strictly positive, strictly decreasing, and strictly convex, assuming the input time series is a scalar-valued functional of a reversible Markov chain. All of the MCMC done by this package is reversible. This **R** function estimates the “big gamma” function, Γ_k considered as a function of k , subject to three different constraints, (1) nonnegative, (2) nonnegative and nonincreasing, and (3) nonnegative, nonincreasing, and convex. It also estimates the variance in the Markov chain central limit theorem (CLT)

$$\gamma_0 + 2 \sum_{k=1}^{\infty} \gamma_k = -\gamma_0 + 2 \sum_{k=0}^{\infty} \Gamma_k$$

Note: The batch means provided by **metrop** are also scalar functionals of a reversible Markov chain. Thus these initial sequence estimators applied to the batch means give valid standard errors for the mean of the batch means even when the batch length is too short to provide a valid estimate of asymptotic variance. One does, of course, have to multiply the asymptotic variance of the batch means by the batch length to get the asymptotic variance for the unbatched chain.

Value

a list containing the following components:

<code>gamma0</code>	the scalar γ_0 , the marginal variance of <code>x</code> .
<code>Gamma.pos</code>	the vector Γ , estimated so as to be nonnegative, where, as always, R uses one-origin indexing so <code>Gamma.pos[1]</code> is Γ_0 .
<code>Gamma.dec</code>	the vector Γ , estimated so as to be nonnegative and nonincreasing, where, as always, R uses one-origin indexing so <code>Gamma.dec[1]</code> is Γ_0 .
<code>Gamma.con</code>	the vector Γ , estimated so as to be nonnegative and nonincreasing and convex, where, as always, R uses one-origin indexing so <code>Gamma.con[1]</code> is Γ_0 .
<code>var.pos</code>	the scalar $-\text{gamma0} + 2 * \text{sum}(\text{Gamma.pos})$, which is the asymptotic variance in the Markov chain CLT. Divide by <code>length(x)</code> to get the approximate variance of the sample mean of <code>x</code> .
<code>var.dec</code>	the scalar $-\text{gamma0} + 2 * \text{sum}(\text{Gamma.dec})$, which is the asymptotic variance in the Markov chain CLT. Divide by <code>length(x)</code> to get the approximate variance of the sample mean of <code>x</code> .
<code>var.con</code>	the scalar $-\text{gamma0} + 2 * \text{sum}(\text{Gamma.con})$, which is the asymptotic variance in the Markov chain CLT. Divide by <code>length(x)</code> to get the approximate variance of the sample mean of <code>x</code> .

Bugs

Not precisely a bug, but `var.pos`, `var.dec`, and `var.con` can be negative. This happens only when the chain is way too short to estimate the variance, and even then rarely. But it does happen.

References

Geyer, C. J. (1992) Practical Markov Chain Monte Carlo. *Statistical Science* 7 473–483.

See Also

[metrop](#)

Examples

```
n <- 2e4
rho <- 0.99
x <- arima.sim(model = list(ar = rho), n = n)
out <- initseq(x)
## Not run:
plot(seq(along = out$Gamma.pos) - 1, out$Gamma.pos,
      xlab = "k", ylab = expression(Gamma[k]), type = "l")
lines(seq(along = out$Gamma.dec) - 1, out$Gamma.dec, col = "red")
lines(seq(along = out$Gamma.con) - 1, out$Gamma.con, col = "blue")

## End(Not run)
# asymptotic 95% confidence interval for mean of x
mean(x) + c(-1, 1) * qnorm(0.975) * sqrt(out$var.con / length(x))
# estimated asymptotic variance
out$var.con
# theoretical asymptotic variance
(1 + rho) / (1 - rho) * 1 / (1 - rho^2)
# illustrating use with batch means
bm <- apply(matrix(x, nrow = 5), 2, mean)
initseq(bm)$var.con * 5
```

logit

Simulated logistic regression data.

Description

Like it says

Usage

`data(logit)`

Format

A data frame with variables

x1 quantitative predictor.

x2 quantitative predictor.

x3 quantitative predictor.

x4 quantitative predictor.

y Bernoulli response.

Examples

```
library(mcmc)
data(logit)
out <- glm(y ~ x1 + x2 + x3 + x4, family = binomial, data = logit)
summary(out)
```

metrop

Metropolis Algorithm

Description

Markov chain Monte Carlo for continuous random vector using a Metropolis algorithm.

Usage

```
metrop(obj, initial, nbatch, blen = 1, nspac = 1, scale = 1, outfun,
       debug = FALSE, ...)
## S3 method for class 'function'
metrop(obj, initial, nbatch, blen = 1, nspac = 1,
       scale = 1, outfun, debug = FALSE, ...)
## S3 method for class 'metropolis'
metrop(obj, initial, nbatch, blen = 1, nspac = 1,
       scale = 1, outfun, debug = FALSE, ...)
```

Arguments

obj Either an R function or an object of class "metropolis" from a previous invocation of this function.

If a function, it evaluates the log unnormalized probability density of the desired equilibrium distribution of the Markov chain. Its first argument is the state vector of the Markov chain. Other arguments arbitrary and taken from the ... arguments of this function. It should return $-\text{Inf}$ for points of the state space having probability zero under the desired equilibrium distribution. See also Details and Warning.

If an object of class "metropolis", any missing arguments (including the log unnormalized density function) are taken from this object. Also *initial* is ignored and the initial state of the Markov chain is the final state from the run recorded in *obj*.

<code>initial</code>	a real vector, the initial state of the Markov chain. Must be feasible, see Details. Ignored if <code>obj</code> is of class "metropolis".
<code>nbatch</code>	the number of batches.
<code>blen</code>	the length of batches.
<code>nspac</code>	the spacing of iterations that contribute to batches.
<code>scale</code>	controls the proposal step size. If scalar or vector, the proposal is $x + \text{scale} * z$ where x is the current state and z is a standard normal random vector. If matrix, the proposal is $x + \text{scale} \%*\% z$.
<code>outfun</code>	controls the output. If a function, then the batch means of <code>outfun(state, ...)</code> are returned. If a numeric or logical vector, then the batch means of <code>state[outfun]</code> (if this makes sense). If missing, the the batch means of <code>state</code> are returned.
<code>debug</code>	if TRUE extra output useful for testing.
<code>...</code>	additional arguments for <code>obj</code> or <code>outfun</code> .

Details

Runs a “random-walk” Metropolis algorithm, terminology introduced by Tierney (1994), with multivariate normal proposal producing a Markov chain with equilibrium distribution having a specified unnormalized density. Distribution must be continuous. Support of the distribution is the support of the density specified by argument `obj`. The initial state must satisfy `obj(state, ...) > -Inf`. Description of a complete MCMC analysis (Bayesian logistic regression) using this function can be found in the vignette `vignette("demo", "mcmc")`.

Suppose the function coded by the log unnormalized function (either `obj` or `obj$lud`) is actually a log unnormalized density, that is, if w denotes that function, then e^w integrates to some value strictly between zero and infinity. Then the `metrop` function always simulates a reversible, Harris ergodic Markov chain having the equilibrium distribution with this log unnormalized density. The chain is not guaranteed to be geometrically ergodic. In fact it cannot be geometrically ergodic if the tails of the log unnormalized density are sufficiently heavy. The `morph.metrop` function deals with this situation.

Value

an object of class "mcmc", subclass "metropolis", which is a list containing at least the following components:

<code>accept</code>	fraction of Metropolis proposals accepted.
<code>batch</code>	<code>nbatch</code> by <code>p</code> matrix, the batch means, where <code>p</code> is the dimension of the result of <code>outfun</code> if <code>outfun</code> is a function, otherwise the dimension of <code>state[outfun]</code> if that makes sense, and the dimension of <code>state</code> when <code>outfun</code> is missing.
<code>accept.batch</code>	a vector of length <code>nbatch</code> , the batch means of the acceptances.
<code>initial</code>	value of argument <code>initial</code> .
<code>final</code>	final state of Markov chain.
<code>initial.seed</code>	value of <code>.Random.seed</code> before the run.
<code>final.seed</code>	value of <code>.Random.seed</code> after the run.
<code>time</code>	running time of Markov chain from <code>system.time()</code> .

lud	the function used to calculate log unnormalized density, either <code>obj</code> or <code>obj\$lud</code> from a previous run.
nbatch	the argument <code>nbatch</code> or <code>obj\$nbatch</code> .
blen	the argument <code>blen</code> or <code>obj\$blen</code> .
nspac	the argument <code>nspac</code> or <code>obj\$nspac</code> .
outfun	the argument <code>outfun</code> or <code>obj\$outfun</code> .

Description of additional output when `debug = TRUE` can be found in the vignette `debug (./doc/debug.pdf)`.

Warning

If `outfun` is missing or not a function, then the log unnormalized density can be defined without a `...` argument and that works fine. One can define it starting `ludfun <- function(state)` and that works or `ludfun <- function(state, foo, bar)`, where `foo` and `bar` are supplied as additional arguments to `metrop`.

If `outfun` is a function, then both it and the log unnormalized density function can be defined without `...` arguments *if they have exactly the same arguments list* and that works fine. Otherwise it doesn't work. Define these functions by

```
ludfun <- function(state, foo)
outfun <- function(state, bar)
```

and you get an error about unused arguments. Instead define these functions by

```
ludfun <- function(state, foo, ...)
outfun <- function(state, bar, ...)
```

and supply `foo` and `bar` as additional arguments to `metrop`, and that works fine.

In short, the log unnormalized density function and `outfun` need to have `...` in their arguments list to be safe. Sometimes it works when `...` is left out and sometimes it doesn't.

Of course, one can avoid this whole issue by always defining the log unnormalized density function and `outfun` to have only one argument `state` and use global variables (objects in the R global environment) to specify any other information these functions need to use. That too follows the R way. But some people consider that bad programming practice.

A third option is to define either or both of these functions using a function factory. This is demonstrated in the vignette for this package named `demo`, which is shown by `vignette("demo", "mcmc")`.

Philosophy of MCMC

This function follows the philosophy of MCMC explained the introductory chapter of the *Handbook of Markov Chain Monte Carlo* (Geyer, 2011).

This function automatically does batch means in order to reduce the size of output and to enable easy calculation of Monte Carlo standard errors (MCSE), which measure error due to the Monte Carlo sampling (not error due to statistical sampling — MCSE gets smaller when you run the computer longer, but statistical sampling variability only gets smaller when you get a larger data set). All of

this is explained in the package vignette `vignette("demo", "mcmc")` and in Section 1.10 of Geyer (2011).

This function does not apparently do “burn-in” because this concept does not actually help with MCMC (Geyer, 2011, Section 1.11.4) but the re-entrant property of this function does allow one to do “burn-in” if one wants. Assuming `ludfun`, `start.value`, `scale` have been already defined and are, respectively, an R function coding the log unnormalized density of the target distribution, a valid state of the Markov chain, and a useful scale factor,

```
out <- metrop(ludfun, start.value, nbatch = 1, blen = 1e5, scale = scale)
out <- metrop(out, nbatch = 100, blen = 1000)
```

throws away a run of 100 thousand iterations before doing another run of 100 thousand iterations that is actually useful for analysis, for example,

```
apply(out$batch, 2, mean)
apply(out$batch, 2, sd) / sqrt(out$nbatch)
```

give estimates of posterior means and their MCSE assuming the batch length (here 1000) was long enough to contain almost all of the significant autocorrelation (see Geyer, 2011, Section 1.10, for more on MCSE). The re-entrant property of this function (the second run starts where the first one stops) assures that this is really “burn-in”.

The re-entrant property allows one to do very long runs without having to do them in one invocation of this function.

```
out2 <- metrop(out)
out3 <- metrop(out2)
batch <- rbind(out$batch, out2$batch, out3$batch)
```

produces a result as if the first run had been three times as long.

Tuning

The scale argument must be adjusted so that the acceptance rate is not too low or too high to get reasonable performance. The rule of thumb is that the acceptance rate should be about 25%. But this recommendation (Gelman, et al., 1996) is justified by analysis of a toy problem (simulating a spherical multivariate normal distribution) for which MCMC is unnecessary. There is no reason to believe this is optimal for all problems (if it were optimal, a stronger theorem could be proved). Nevertheless, it is clear that at very low acceptance rates the chain makes little progress (because in most iterations it does not move) and that at very high acceptance rates the chain also makes little progress (because unless the log unnormalized density is nearly constant, very high acceptance rates can only be achieved by very small values of scale so the steps the chain takes are also very small).

Even in the Gelman, et al. (1996) result, the optimal rate for spherical multivariate normal depends on dimension. It is 44% for $d = 1$ and 23% for $d = \infty$. Geyer and Thompson (1995) have an example, admittedly for simulated tempering (see [temper](#)) rather than random-walk Metropolis, in which no acceptance rate less than 70% produces an ergodic Markov chain. Thus 25% is merely a rule of thumb. We only know we don’t want too high or too low. Probably 1% or 99% is very inefficient.

References

- Gelman, A., Roberts, G. O., and Gilks, W. R. (1996) Efficient Metropolis jumping rules. In *Bayesian Statistics 5: Proceedings of the Fifth Valencia International Meeting*. Edited by J. M. Bernardo, J. O. Berger, A. P. Dawid, and A. F. M. Smith. Oxford University Press, Oxford, pp. 599–607.
- Geyer, C. J. (2011) Introduction to MCMC. In *Handbook of Markov Chain Monte Carlo*. Edited by S. P. Brooks, A. E. Gelman, G. L. Jones, and X. L. Meng. Chapman & Hall/CRC, Boca Raton, FL, pp. 3–48.
- Geyer, C. J. and Thompson, E. A. (1995) Annealing Markov chain Monte Carlo with applications to ancestral inference. *Journal of the American Statistical Association* **90** 909–920.
- Tierney, L. (1994) Markov chains for exploring posterior distributions (with discussion). *Annals of Statistics* **22** 1701–1762.

See Also

[morph.metrop](#) and [temper](#)

Examples

```
h <- function(x) if (all(x >= 0) && sum(x) <= 1) return(1) else return(-Inf)
out <- metrop(h, rep(0, 5), 1000)
out$accept
# acceptance rate too low
out <- metrop(out, scale = 0.1)
out$accept
t.test(out$accept.batch)$conf.int
# acceptance rate o. k. (about 25 percent)
plot(out$batch[ , 1])
# but run length too short (few excursions from end to end of range)
out <- metrop(out, nbatch = 1e4)
out$accept
plot(out$batch[ , 1])
hist(out$batch[ , 1])
acf(out$batch[ , 1], lag.max = 250)
# looks like batch length of 250 is perhaps OK
out <- metrop(out, blen = 250, nbatch = 100)
apply(out$batch, 2, mean) # Monte Carlo estimates of means
apply(out$batch, 2, sd) / sqrt(out$nbatch) # Monte Carlo standard errors
t.test(out$accept.batch)$conf.int
acf(out$batch[ , 1]) # appears that blen is long enough
```

morph

Variable Transformation

Description

Utility functions for variable transformation.

Usage

```
morph(b, r, p, center)
morph.identity()
```

Arguments

b	Positive real number. May be missing.
r	Non-negative real number. May be missing. If p is specified, r defaults to 0.
p	Real number strictly greater than 2. May be missing. If r is specified, p defaults to 3.
center	Real scalar or vector. May be missing. If center is a vector it should be the same length of the state of the Markov chain, center defaults to 0

Details

The morph function facilitates using variable transformations by providing functions to (using X for the original random variable with the pdf f_X , and Y for the transformed random variable with the pdf f_Y):

- Calculate the log unnormalized probability density for Y induced by the transformation.
- Transform an arbitrary function of X to a function of Y .
- Transform values of X to values of Y .
- Transform values of Y to values of X (the inverse transformation).

for a select few transformations.

`morph.identity` implements the identity transformation, $Y = X$.

The parameters r , p , b and `center` specify the transformation function. In all cases, `center` gives the center of the transformation, which is the value c in the equation

$$Y = f(X - c).$$

If no parameters are specified, the identity transformation, $Y = X$, is used.

The parameters r , p and b specify a function g , which is a monotonically increasing bijection from the non-negative reals to the non-negative reals. Then

$$f(X) = g(|X|) \frac{X}{|X|}$$

where $|X|$ represents the Euclidean norm of the vector X . The inverse function is given by

$$f^{-1}(Y) = g^{-1}(|Y|) \frac{Y}{|Y|}.$$

The parameters r and p are used to define the function

$$g_1(x) = x + (x - r)^p I(x > r)$$

where $I(\cdot)$ is the indicator function. We require that r is non-negative and p is strictly greater than 2. The parameter b is used to define the function

$$g_2(x) = (e^{bx} - e/3)I(x > \frac{1}{b}) + (x^3b^3e/6 + xbe/2)I(x \leq \frac{1}{b})$$

We require that b is positive.

The parameters r , p and b specify f^{-1} in the following manner:

- If one or both of r and p is specified, and b is not specified, then

$$f^{-1}(X) = g_1(|X|) \frac{X}{|X|}.$$

If only r is specified, $p = 3$ is used. If only p is specified, $r = \emptyset$ is used.

- If only b is specified, then

$$f^{-1}(X) = g_2(|X|) \frac{X}{|X|}.$$

- If one or both of r and p is specified, and b is also specified, then

$$f^{-1}(X) = g_2(g_1(|X|)) \frac{X}{|X|}.$$

Value

a list containing the functions

- `outfun(f)`, a function that operates on functions. `outfun(f)` returns the function `function(state, ...) f(inverse(state), ...)`.
- `inverse`, the inverse transformation function.
- `transform`, the transformation function.
- `lud`, a function that operates on functions. As input, `lud` takes a function that calculates a log unnormalized probability density, and returns a function that calculates the log unnormalized density by transforming a random variable using the transform function. `lud(f) = function(state, ...) f(inverse(state), ...) + log.jacobian(state)`, where `log.jacobian` represents the function that calculate the log Jacobian of the transformation. `log.jacobian` is not returned.

Warning

The equations for the returned transform function (see below) do not have a general analytical solution when p is not equal to 3. This implementation uses numerical approximation to calculate transform when p is not equal to 3. If computation speed is a factor, it is advisable to use $p=3$. This is not a factor when using [morph.metrop](#), as `transform` is only called once during setup, and not at all while running the Markov chain.

See Also

[morph.metrop](#)

Examples

```
# use an exponential transformation, centered at 100.
b1 <- morph(b=1, center=100)
# original log unnormalized density is from a t distribution with 3
# degrees of freedom, centered at 100.
lud.transformed <- b1$lud(function(x) dt(x - 100, df=3, log=TRUE))
d.transformed <- Vectorize(function(x) exp(lud.transformed(x)))
## Not run:
curve(d.transformed, from=-3, to=3, ylab="Induced Density")

## End(Not run)
```

morph.metrop

Morphometric Metropolis Algorithm

Description

Markov chain Monte Carlo for continuous random vector using a Metropolis algorithm for an induced density.

Usage

```
morph.metrop(obj, initial, nbatch, blen = 1, nspac = 1, scale = 1,
  outfun, debug = FALSE, morph, ...)
```

Arguments

obj	see metrop .
initial	see metrop .
nbatch	see metrop .
blen	see metrop .
nspac	see metrop .
scale	see metrop .
outfun	unlike for metrop must be a function or missing; if missing the identity function, <code>function(x) x</code> , is used.
debug	see metrop .
morph	morph object used for transformations. See morph .
...	see metrop .

Details

morph.metrop implements morphometric methods for Markov chains. The caller specifies a log unnormalized probability density and a transformation. The transformation specified by the morph parameter is used to induce a new log unnormalized probability density, a Metropolis algorithm is run for the induced density. The Markov chain is transformed back to the original scale. Running the Metropolis algorithm for the induced density, instead of the original density, can result in a Markov chain with better convergence properties. For more details see Johnson and Geyer (submitted). Except for morph, all parameters are passed to `metrop`, transformed when necessary. The scale parameter is *not* transformed.

If X is a real vector valued continuous random variable, and $Y = f(X)$ where f is a diffeomorphism, then the pdf of Y is given by

$$f_Y(y) = f_X(f^{-1}(y))|\nabla f^{-1}(y)|$$

where f_X is the pdf of X and ∇f^{-1} is the Jacobian of f^{-1} . Because f is a diffeomorphism, a Markov chain for f_Y may be transformed into a Markov chain for f_X . Furthermore, these Markov chains are isomorphic (Johnson and Geyer, submitted) and have the same convergence rate. The `morph` variable provides a diffeomorphism, `morph.metrop` uses this diffeomorphism to induce the log unnormalized density, $\log f_Y$ based on the user supplied log unnormalized density, $\log f_X$. `morph.metrop` runs a Metropolis algorithm for $\log f_Y$ and transforms the resulting Markov chain into a Markov chain for f_X . The user accessible output components are the same as those that come from `metrop`, see the documentation for `metrop` for details.

Subsequent calls of `morph.metrop` may change to the transformation by specifying a new value for `morph`.

Any of the other parameters to `morph.metrop` may also be modified in subsequent calls. See `metrop` for more details.

The general idea is that a random-walk Metropolis sampler (what `metrop` does) will not be geometrically ergodic unless the tails of the unnormalized density decrease superexponentially fast (so the tails of the log unnormalized density decrease faster than linearly). It may not be geometrically ergodic even then (see Johnson and Geyer, submitted, for the complete theory). The transformations used by this function (provided by `morph`) can produce geometrically ergodic chains when the tails of the log unnormalized density are too light for `metrop` to do so.

When the tails of the unnormalized density are exponentially light but not superexponentially light (so the tails of the log unnormalized density are asymptotically linear, as in the case of exponential family models when conjugate priors are used, for example logistic regression, Poisson regression with log link, or log-linear models for categorical data), one should use `morph` with `b = 0` (the default), which produces a transformation of the form g_1 in the notation used in the details section of the help for `morph`. This will produce a geometrically ergodic sampler if other features of the log unnormalized density are well behaved. For example it will do so for the exponential family examples mentioned above. (See Johnson and Geyer, submitted, for the complete theory.)

The transformation g_1 behaves like a shift transformation on a ball of radius r centered at center, so these arguments to `morph` should be chosen so that a sizable proportion of the probability under the original (untransformed) unnormalized density is contained in this ball. This function will work when `r = 0` and `center = 0` (the defaults) are used, but may not work as well as when r and center are well chosen.

When the tails of the unnormalized density are not exponentially light (so the tails of the log unnormalized density decrease sublinearly, as in the case of univariate and multivariate t distributions),

one should use `morph` with $r > 0$ and $p = 3$, which produces a transformation of the form g_2 composed with g_1 in the notation used in the details section of the help for `morph`. This will produce a geometrically ergodic sampler if other features of the log unnormalized density are well behaved. For example it will do so for the t examples mentioned above. (See Johnson and Geyer, submitted, for the complete theory.)

Value

an object of class `mcmc`, subclass `morph.metropolis`. This object is a list containing all of the elements from an object returned by `metrop`, plus at least the following components:

`morph` the morph object used for the transformations.
`morph.final` the final state of the Markov chain on the transformed scale.

References

Johnson, L. T. and Geyer, C. J. (submitted) Variable Transformation to Obtain Geometric Ergodicity in the Random-walk Metropolis Algorithm.

See Also

`metrop`, `morph`.

Examples

```
out <- morph.metrop(function(x) dt(x, df=3, log=TRUE), 0, blen=100,
  nbatch=100, morph=morph(b=1))
# change the transformation.
out <- morph.metrop(out, morph=morph(b=2))
out$accept
# accept rate is high, increase the scale.
out <- morph.metrop(out, scale=4)
# close to 0.20 is about right.
out$accept
```

olbm

Overlapping Batch Means

Description

Variance of sample mean of time series calculated using overlapping batch means.

Usage

```
olbm(x, batch.length, demean = TRUE)
```

Arguments

<code>x</code>	a matrix or time series object. Each column of <code>x</code> is treated as a scalar time series.
<code>batch.length</code>	length of batches.
<code>demean</code>	when <code>demean = TRUE</code> (the default) the sample mean is subtracted from each batch mean when estimating the variance. Using <code>demean = FALSE</code> would essentially assume the true mean is known to be zero, which might be useful in a toy problem where the answer is known.

Value

The estimated variance of the sample mean.

See Also

[ts](#)

Examples

```
h <- function(x) if (all(x >= 0) && sum(x) <= 1) return(1) else return(-Inf)
out <- metrop(h, rep(0, 5), 1000)
out <- metrop(out, scale = 0.1)
out <- metrop(out, nbatch = 1e4)
foo <- olbm(out$batch, 150)
# monte carlo estimates (true means are same by symmetry)
apply(out$batch, 2, mean)
# monte carlo standard errors (true s. d. are same by symmetry)
sqrt(diag(foo))
# check that batch length is reasonable
acf(out$batch, lag.max = 200)
```

temper

Simulated Tempering and Umbrella Sampling

Description

Markov chain Monte Carlo (MCMC) for continuous random vectors using parallel or serial tempering, the latter also called umbrella sampling and simulated tempering. The chain simulates k different distributions on the same state space. In parallel tempering, all the distributions are simulated in each iteration. In serial tempering, only one of the distributions is simulated (a random one). In parallel tempering, the state is a $k \times p$ matrix, each row of which is the state for one of the distributions. In serial tempering the state of the Markov chain is a pair (i, x) , where i is an integer between 1 and k and x is a vector of length p . This pair is represented as a single real vector $c(i, x)$. The variable i says which distribution x is a simulation for.

Usage

```

temper(obj, initial, neighbors, nbatch, blen = 1, nspac = 1, scale = 1,
       outfun, debug = FALSE, parallel = FALSE, ...)
## S3 method for class 'function'
temper(obj, initial, neighbors, nbatch,
       blen = 1, nspac = 1, scale = 1,
       outfun, debug = FALSE, parallel = FALSE, ...)
## S3 method for class 'tempering'
temper(obj, initial, neighbors, nbatch,
       blen = 1, nspac = 1, scale = 1,
       outfun, debug = FALSE, parallel = FALSE, ...)

```

Arguments

obj	<p>either an R function or an object of class "tempering" from a previous run.</p> <p>If a function, it should evaluate the log unnormalized density $\log h(i, x)$ of the desired equilibrium distribution of the Markov chain for serial tempering (the same function is used for both serial and parallel tempering, see Details below for further explanation).</p> <p>If an object of class "tempering", the log unnormalized density function is <code>obj\$lud</code>, and missing arguments of <code>temper</code> are obtained from the corresponding elements of <code>obj</code>.</p> <p>The first argument of the log unnormalized density function is the is an R vector $c(i, x)$, where i says which distribution x is supposed to be a simulation for. Other arguments are arbitrary and taken from the <code>...</code> arguments of <code>temper</code>. The log unnormalized density function should return <code>-Inf</code> in regions of the state space having probability zero.</p>
initial	<p>for serial tempering, a real vector $c(i, x)$ as described above. For parallel tempering, a real $k \times p$ matrix as described above. In either case, the initial state of the Markov chain. Ignored if <code>obj</code> has class "tempering".</p>
neighbors	<p>a logical symmetric matrix of dimension k by k. Elements that are TRUE indicate jumps or swaps attempted by the Markov chain. Ignored if <code>obj</code> has class "tempering".</p>
nbatch	<p>the number of batches.</p>
blen	<p>the length of batches.</p>
nspac	<p>the spacing of iterations that contribute to batches.</p>
scale	<p>controls the proposal step size for real elements of the state vector. For serial tempering, proposing a new value for the x part of the state (i, x). For parallel tempering, proposing a new value for the x_i part of the state (x_1, \dots, x_k). In either case, the proposal is a real vector of length p. If scalar or vector, the proposal is $x + \text{scale} * z$ where x is the part x or x_i of the state the proposal may replace. If matrix, the proposal is $x + \text{scale} \%*\% z$. If list, the length must be k, and each element must be scalar, vector, or matrix, and operate as described above. The i-th component of the list is used to update x when the state is (i, x) or x_i otherwise.</p>

outfun	controls the output. If a function, then the batch means of <code>outfun(state, ...)</code> are returned. The argument <code>state</code> is like the argument <code>initial</code> of this function. If missing, the batch means of the real part of the state vector or matrix are returned, and for serial tempering the batch means of a multivariate Bernoulli indicating the current component are returned.
debug	if TRUE extra output useful for testing.
parallel	if TRUE does parallel tempering, if FALSE does serial tempering. Ignored if <code>obj</code> has class "tempering".
...	additional arguments for <code>obj</code> or <code>outfun</code> .

Details

Serial tempering simulates a mixture of distributions of a continuous random vector. The number of components of the mixture is k , and the dimension of the random vector is p . Denote the state (i, x) , where i is a positive integer between 1 and k , and let $h(i, x)$ denote the unnormalized joint density of their equilibrium distribution. The logarithm of this function is what `obj` or `obj$lud` calculates. The mixture distribution is the marginal for x derived from the equilibrium distribution $h(i, x)$, that is,

$$h(x) = \sum_{i=1}^k h(i, x)$$

Parallel tempering simulates a product of distributions of a continuous random vector. Denote the state (x_1, \dots, x_k) , then the unnormalized joint density of the equilibrium distribution is

$$h(x_1, \dots, x_k) = \prod_{i=1}^k h(i, x_i)$$

The update mechanism of the Markov chain combines two kinds of elementary updates: jump/swap updates (jump for serial tempering, swap for parallel tempering) and within-component updates. Each iteration of the Markov chain one of these elementary updates is done. With probability 1/2 a jump/swap update is done, and with probability 1/2 a with-component update is done.

Within-component updates are the same for both serial and parallel tempering. They are "random-walk" Metropolis updates with multivariate normal proposal, the proposal distribution being determined by the argument `scale`. In serial tempering, the x part of the current state (i, x) is updated preserving $h(i, x)$. In parallel tempering, an index i is chosen at random and the part of the state x_i representing that component is updated, again preserving $h(i, x)$.

Jump updates choose uniformly at random a neighbor of the current component: if i indexes the current component, then it chooses uniformly at random a j such that `neighbors[i, j] == TRUE`. It then does a Metropolis-Hastings update for changing the current state from (i, x) to (j, x) .

Swap updates choose a component uniformly at random and a neighbor of that component uniformly at random: first an index i is chosen uniformly at random between 1 and k , then an index j is chosen uniformly at random such that `neighbors[i, j] == TRUE`. It then does a Metropolis-Hastings update for swapping the states of the two components: interchanging x_i and x_j while preserving $h(x_1, \dots, x_k)$.

The initial state must satisfy `lud(initial, ...)` > - Inf for serial tempering or must satisfy `lud(initial[i,], ...)` > - Inf for each i for parallel tempering, where `lud` is either `obj` or `obj$lud`. That is, the initial state must have positive probability.

Value

an object of class "mcmc", subclass "tempering", which is a list containing at least the following components:

batch	the batch means of the continuous part of the state. If outfun is missing, an nbatch by k by p array. Otherwise, an nbatch by m matrix, where m is the length of the result of outfun.
ibatch	(returned for serial tempering only) an nbatch by k matrix giving batch means for the multivariate Bernoulli random vector that is all zeros except for a 1 in the i-th place when the current state is (i, x) .
acceptx	fraction of Metropolis within-component proposals accepted. A vector of length k giving the acceptance rate for each component.
accepti	fraction of Metropolis jump/swap proposals accepted. A k by k matrix giving the acceptance rate for each allowed jump or swap component. NA for elements such that the corresponding elements of neighbors is FALSE.
initial	value of argument initial.
final	final state of Markov chain.
initial.seed	value of .Random.seed before the run.
final.seed	value of .Random.seed after the run.
time	running time of Markov chain from system.time().
lud	the function used to calculate log unnormalized density, either obj or obj\$lud from a previous run.
nbatch	the argument nbatch or obj\$nbatch.
blen	the argument blen or obj\$blen.
nspac	the argument nspac or obj\$nspac.
outfun	the argument outfun or obj\$outfun.

Description of additional output when debug = TRUE can be found in the vignette debug, which is shown by vignette("debug", "mcmc").

Warning

If outfun is missing, then the log unnormalized density function can be defined without a ... argument and that works fine. One can define it starting ludfun <- function(state) and that works or ludfun <- function(state, foo, bar), where foo and bar are supplied as additional arguments to temper and that works too.

If outfun is a function, then both it and the log unnormalized density function can be defined without ... arguments *if they have exactly the same arguments list* and that works fine. Otherwise it doesn't work. Define these functions by

```
ludfun <- function(state, foo)
outfun <- function(state, bar)
```

and you get an error about unused arguments. Instead define these functions by

```
ludfun <- function(state, foo, ...)
outfun <- function(state, bar, ...)
```

and supply `foo` and `bar` as additional arguments to `temper`, and that works fine.

In short, the log unnormalized density function and `outfun` need to have `...` in their arguments list to be safe. Sometimes it works when `...` is left out and sometimes it doesn't.

Of course, one can avoid this whole issue by always defining the log unnormalized density function and `outfun` to have only one argument `state` and use global variables (objects in the R global environment) to specify any other information these functions need to use. That too follows the R way. But some people consider that bad programming practice.

A third option is to define either or both of these functions using a function factory. This is demonstrated in the vignette for this package named `demo`, which is shown by `vignette("demo", "mcmc")`.

Philosophy of MCMC

This function follows the philosophy of MCMC explained in the introductory chapter of the *Handbook of Markov Chain Monte Carlo* (Geyer, 2011a) and in the chapter of that book on tempering and related subjects (Geyer, 2011b). See also the section on philosophy of `metrop`.

Tuning

The `scale` argument must be adjusted so that the acceptance rate for within-component proposals (component `acceptx` of the result returned by this function) is not too low or too high to get reasonable performance. The log unnormalized density function must be chosen so that the acceptance rate for jump/swap proposals (component `accepti` of the result returned by this function) is not too low or too high to get reasonable performance. The former is a vector and the latter is a matrix, and all these rates must be adjusted to be reasonable.

The rates in `accepti` are influenced by the number of components of the tempering mixture distribution, by what those components are (how far apart they are in some unspecified metric on probability distributions), and by the chosen normalizing constants for those distributions.

For examples of tuning tempering, see Geyer (2011b) and also the vignette of this package shown by `vignette("bfst", "mcmc")`. The help for R function `metrop` also gives advice on tuning its sampler, which is relevant for tuning the `acceptx` rates.

See also Geyer (1991) and Geyer and Thompson (1995) for the general theory of tuning parallel and serial tempering.

References

- Geyer, C. J. (1991) Markov chain Monte Carlo maximum likelihood. *Computing Science and Statistics: Proc. 23rd Symp. Interface*, 156–163. <http://hdl.handle.net/11299/58440>.
- Geyer, C. J. (2011a) Introduction to MCMC. In *Handbook of Markov Chain Monte Carlo*. Edited by S. P. Brooks, A. E. Gelman, G. L. Jones, and X. L. Meng. Chapman & Hall/CRC, Boca Raton, FL, pp. 3–48.
- Geyer, C. J. (2011b) Importance Sampling, Simulated Tempering, and Umbrella Sampling. In *Handbook of Markov Chain Monte Carlo*. Edited by S. P. Brooks, A. E. Gelman, G. L. Jones, and X. L. Meng. Chapman & Hall/CRC, Boca Raton, FL, pp. 295–312.

Geyer, C. J. and Thompson, E. A. (1995) Annealing Markov chain Monte Carlo with applications to ancestral inference. *Journal of the American Statistical Association* **90** 909–920.

Examples

```
d <- 9
witch.which <- c(0.1, 0.3, 0.5, 0.7, 1.0)
ncomp <- length(witch.which)

neighbors <- matrix(FALSE, ncomp, ncomp)
neighbors[row(neighbors) == col(neighbors) + 1] <- TRUE
neighbors[row(neighbors) == col(neighbors) - 1] <- TRUE

ludfun <- function(state, log.pseudo.prior = rep(0, ncomp)) {
  stopifnot(is.numeric(state))
  stopifnot(length(state) == d + 1)
  icomp <- state[1]
  stopifnot(icomp == as.integer(icomp))
  stopifnot(1 <= icomp && icomp <= ncomp)
  stopifnot(is.numeric(log.pseudo.prior))
  stopifnot(length(log.pseudo.prior) == ncomp)
  theta <- state[-1]
  if (any(theta > 1.0)) return(-Inf)
  bnd <- witch.which[icomp]
  lpp <- log.pseudo.prior[icomp]
  if (any(theta > bnd)) return(lpp)
  return(- d * log(bnd) + lpp)
}

# parallel tempering
thetas <- matrix(0.5, ncomp, d)
out <- temper(ludfun, initial = thetas, neighbors = neighbors, nbatch = 20,
  blen = 10, nspac = 5, scale = 0.56789, parallel = TRUE, debug = TRUE)

# serial tempering
theta.initial <- c(1, rep(0.5, d))
# log pseudo prior found by trial and error
qux <- c(0, 9.179, 13.73, 16.71, 20.56)

out <- temper(ludfun, initial = theta.initial, neighbors = neighbors,
  nbatch = 50, blen = 30, nspac = 2, scale = 0.56789,
  parallel = FALSE, debug = FALSE, log.pseudo.prior = qux)
```

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